Partial Identification of Dose Responses with Hidden Confounders

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Confounders?

Wine May Help With Diabetes

Drink to Your Health (in Moderation), the Science Says

2023

2015

2015

2008

2008

2008

2004

2002

1994

New Hints Seen That Red Wine May Slow Aging

Red Wine May Curb Fat Cells

Regimens: Wine May Help Keep Liver Healthy

Alcohol's Benefits Extend to Hypertension

Evidence Mounting That Moderate Drinking Is Healthful

Even a Little Alcohol Can Harm Your Health

Recent research makes it clear that any amount of drinking can be detrimental. Here's why you may want to cut down on your consumption beyond Dry January.

> Wine for the Heart: Over All, Risks May Outweigh Benefits





Why Partial Identification [...] with Hidden Confounders

- Typical predictions are descriptive.
- Causal inferences are prescriptive.
 - We aim to predict the outcome of an intervention.
 - If a study produces *actionable* insights, then it is claiming to make a causal inference, whether explicitly or not.
- If we know there might be confounders (endogeneity), then point identification of causal outcomes is impossible.
- Partial identification is our best bet: produce a set of outcomes admitted by the causal setting.

Why Dose Responses

- "Amount" of treatment is important in many problems.
- Derivative of the curve gives incremental effects.



Related: Simpson's Paradox



Our Proposed Method

Potential Outcomes Setup

Say we have an outcome prediction model. Assume we learned it perfectly.

outcon

- This predicts the potential outcome at
- everything else that is relevant in the problem, i.e. confounders.
- The dose response is $\mathbb{E}[Y_t \mid X = x]$ as a function of t.

ne covariates
$$\int y|t, x)$$

treatment

t.
$$p(y_t|t, x) = p(y|t, x)$$

• "Potential outcomes" are the different treatment outcomes after controlling for

The Ignorability Assumption

• No hidden confounding!



Hidden Confounding

- If the ignorability assumption held, then assigned treatment wouldn't affect the potential outcome, conditioned on observed confounders.
- In that case,

$$p(y_t|x) = p(y_t|t, x) = p(y|t, x)$$

- However, a hidden confounder could *ruin* this via a backdoor path.
- The graph to the right gives one such example with the red arrows.



Continuous Treatments

 $p(y_t|x) = \int_{\mathcal{T}} p(y_t|\tau, x) p(\tau|x) \, \mathrm{d}\tau$

the potential outcome

Continuous Treatments

 $p(y_t|x) = \int_{\mathcal{T}} p(y_t|\tau, x) p(\tau|x) \, \mathrm{d}\tau$ potential outcome the potential outcome $p(y_t \mid \tau, x)$

"what is this person's potential outcome at t given that their assigned treatment is tau"

the counterfactual

The Problem with Continuous Treatments

a) Confounded Outcomes for <u>Binary</u> Treatments

$P[Y_{t}] = P[Y_{t}|T = t] \times P[T = t]$ + $P[Y_t | T = 1-t] \times P[T = 1-t]$ counterfactual

The Problem with **Continuous Treatments**

a) Confounded Outcomes for <u>Binary</u> Treatments

b) Confounded Outcomes for <u>Continuous</u> Treatments

- Infinite unobservable counterfactuals!
- The integrand cannot be identified almost anywhere.





The Problem with **Continuous Treatments**

a) Confounded Outcomes for <u>Binary</u> Treatments

b) Confounded Outcomes for Continuous Treatments

- Infinite unobservable counterfactuals!
- The integrand cannot be identified almost anywhere.
- We need an approximation.





We Know Nothing!

• First step to the solution is extrapolating from the point that we can observe.



$$+ (\tau - t)\partial_{\tau} p(y_t | \tau, x)|_{\tau = t}$$

$$p(y_t | \tau, x)|_{\tau = t} + \mathcal{O}(\tau - t)^3 e^{x t^{rapola}}$$



Now We Know Something

(from before)

 $p(y_t|x) = \int_{\mathcal{T}} p(y_t|\tau, x) p(\tau|x) \,\mathrm{d}\tau$

Second step is to specify where that extrapolation can be trusted, and how much.

Now We Know Something

(from before)

 $p(y_t|x) =$

(use weights and split)

Second step is to specify where that extrapolation can be trusted, and how much.

$$= \int_{\mathcal{T}} p(y_t | \tau, x) p(\tau | x) \, \mathrm{d}\tau$$

$$\approx \int_{\mathcal{T}} \underbrace{w_t(\tau) \tilde{p}(y_t | \tau, x) p(\tau | x) \, \mathrm{d}\tau}_{(A) \text{ the approximated quantity}}$$

$$+ \int_{\mathcal{T}} \underbrace{[1 - w_t(\tau)] p(\tau | y_t, x) p(y_t | x) \, \mathrm{d}\tau}_{(B) \text{ by Bayes' rule}}$$

Solution Outline

We find that

 $p(y_t|x) \approx \frac{\int_{\mathcal{T}} w_t(\tau)}{\int_{\mathcal{T}} w_t}$

- We'll figure out what to do with the "trust weights" later.
- We still have two unknowns: the approximation, and the denominator.

$$-) \widetilde{p}(y_t | \tau, x) p(\tau | x) \,\mathrm{d} au$$

 $\overline{v_t(\tau) p(\tau | y_t, x) \,\mathrm{d} au}$.

Finally Introducing the Sensitivity Model, δMSM

the two counterfactuals, i.e. the ratio of their densities.

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- For binary treatments, the classical MSM bounds the Radon-Nikodym derivative of the two counterfactuals, i.e. the ratio of their densities.
- We follow a similar route but take it to the infinitesimal limit:

 $\omega_{\delta}(y_t | \tau, x) :=$

$$= \frac{p(y_t | \tau + \delta, x)}{p(y_t | \tau, x)}$$

ratio of nearby counterfactuals

Finally Introducing the Sensitivity Model, δMSM

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 $\omega_{\delta}(y_t | \tau, x)$:



nominal pro

$$= \frac{p(y_t \mid \tau + \delta, x)}{p(y_t \mid \tau, x)}$$

$$= \frac{\delta \mid x}{\left| x \right|} \int_{-1}^{-1} \left[\frac{p(\tau + \delta \mid y_t, x)}{p(\tau \mid y_t, x)} \right] \quad \text{by Baye}$$

$$= \frac{\left[\sum_{t=1}^{n} \frac{p(\tau + \delta \mid y_t, x)}{p(\tau \mid y_t, x)} \right]}{\sum_{t=1}^{n} \frac{p(\tau \mid y_t, x)}{p(\tau \mid y_t, x)}} \quad \text{by Baye}$$



Definition of the δMSM

For treatments $t \in \mathcal{T} \subseteq \mathbb{R}$, where \mathcal{T} is connected, and violation-of-ignorability factor $\Gamma \geq 1$, the δ MSM requires

 $\left|\frac{\partial}{\partial\tau}\log\frac{p(\tau)}{p(\tau)}\right|$

$$\left| \frac{|y_t, x)}{(\tau | x)} \right| \le \log \Gamma$$

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ratio of complete and nominal propensities

"how much is this person's treatment assignment informed by a potential outcome (through backdoor paths)"

$$\left| \frac{y_t, x}{(\tau | x)} \right| \le \log \Gamma$$

Necessary Assumptions for Hidden Confounding

- First Assumption: δ MSM holds with some Γ .
- Second Assumption: we need an "anchor point," designated as zero treatment.

$$p(\tau = 0 \mid y_t, x) = p(\tau$$

 $= 0 \mid x$) for all x, t, and y_t

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- This is necessary to solve the integrals.
- What does the second assumption mean for our partial identification?
 - Informally, hidden confounders "matter less" at near-zero treatment values.

= 0 | x) for all x, t, and y_t

Combining the Ingredients

partial identification factor

approx. potential outcome

 $\tilde{p}(y_t|x) = d(t|y_t, x)^{-1} p(y_t|t, x)$

outcome prediction

Combining the Ingredients

partial identification factor

 $\tilde{p}(y_t|x) = d(t|$

approx. potential outcome

Admissible probability densities are governed by

$$d(t|y_t, x) \in \left[\underline{d}(t|y_t, x), \overline{d}(t|y_t, x)\right]$$

which can be solved in closed form! Note: d < 1 < d

$$y_t, x)^{-1} p(y_t | t, x)$$

outcome prediction

Settling the Trust Weights

- density with respect to treatment assignment.

(and form) as the nominal propensities, but always centered at t of course.

We found solutions for various exponential families.

Accuracy of the extrapolation depends on the continuity of the counterfactual

• Narrower treatment propensity densities p(au|x) suggest worse extrapolations.

- Therefore, we parametrize the weights $\,w_t(au)$ to have the same narrowness



T≠t

T≈t



One Last Thing: Relaxing the Second Assumption

• For Beta-distributed treatments, we symmetrify the anchor point assumption.

$$\begin{cases} p(\tau = 0 \mid y_t, x) = p\\ p(\tau = 1 \mid y_t, x) = p \end{cases}$$

about treatment assignment.

 $p(\tau = 0 | x)$ w.p. t, $p(\tau = 1 | x)$ w.p. 1 - t

• New interpretation: the more distant the potential outcome, the less informative it is



Relaxed Anchor Points, Illustrated

- Are you the kind of person that drinks a lot of wine? ($\tau = 1$)
 - Depends on your health outcome from drinking a lot of wine. (y_1)
 - Depends on your health outcome from drinking no wine. (y_0)
- Are you the kind of person that drinks no wine? ($\tau=0$)
 - Depends on your health outcome from drinking a lot of wine. (y_1)
 - Depends on your health outcome from drinking no wine. (y_0)

Benchmark Results

Benchmarks	brain		blood		pbmc		mftc		9	ratio
	mean	(std.)	mean	(std.)	mean	(std.)	mean	(std.)	% best	to best
δ MSM (ours)	138	(120)	141	(129)	138	(121)	144	(124)	78.4	$1.03\ (0.08)$
CMSM	186	(153)	188	(156)	205	(169)	182	(145)	7.8	1.81(2.15)
uniform	158	(137)	162	(146)	157	(136)	167	(141)	4.8	1.20(0.10)
binary MSM	211	(128)	213	(131)	222	(127)	214	(127)	9.0	2.57(2.34)

- Partial-identification costs of 90% coverage of the average dose responses.
- Semi-synthetic confounders are random projections of original data.
- Random quadratic forms describe the potential outcome.
- 500 experiments per dataset and method.

Conclusion — So What

- We described the first sensitivity model for continuous treatments that
 - changes with the propensity (& in a sensical way)
 - always admits valid potential outcome densities.

• Extensive semi-synthetic benchmarks show consistently superior performance to baseline sensitivity models.



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