

Belief functions for the working scientist

A UAI 2015 Tutorial

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UAI 2015

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Some history

- a mathematical framework for representing and reasoning with uncertain information
- also known as **Dempster-Shafer (DS) theory** or **Evidence theory**
- originates from the work of Dempster (1968) in the context of **statistical inference**
- formalized by Shafer (1976) as a **theory of evidence**
- popularized and developed by Smets in the 1980's and 1990's under the name **Transferable Belief Model**.
- starting from the 1990's, **growing number of applications** in AI, information fusion, classification, reliability and risk analysis, etc.

Rationale

- 1 a modeling language for representing **elementary items of evidence and combining them**, in order to form a representation of our beliefs about certain aspects of the world
- 2 the theory of belief function subsumes both the **set-based** and **probabilistic** approaches to uncertainty:
 - a belief function may be viewed both as a **generalized set** and as a **non additive measure**
 - basic mechanisms for reasoning with belief functions extend both probabilistic operations (such as marginalization and conditioning) and set-theoretic operations (such as intersection and union)
- 3 DS reasoning produces the same results as probabilistic reasoning or interval analysis when provided with the same information
- 4 however, its **greater expressive power** allows us to handle more general forms of information

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Mass functions

or “Basic Probability Assignments”

- let ω be an unknown quantity with possible values in a **finite** domain Ω , called the **frame of discernment**
- a piece of evidence about ω may be represented by a **mass function** m on Ω , defined as a function $2^\Omega \rightarrow [0, 1]$, such that:

$$m(\emptyset) = 0 \quad \sum_{A \subseteq \Omega} m(A) = 1$$

- $\mathcal{P}(\Omega) = 2^\Omega$ is the set of all subsets of Ω
- any subset A of Ω such that $m(A) > 0$ is called a **focal element** (FE) or “focal set” of m
- special cases:
 - a **logical** (or “categorical”) mass function has one focal set (\sim set)
 - a **Bayesian** mass function has only focal sets of cardinality one (\sim probability distribution)
- **complete ignorance** is represented by the vacuous mass function defined by $m_\Omega(\omega) = 1$

Mass function

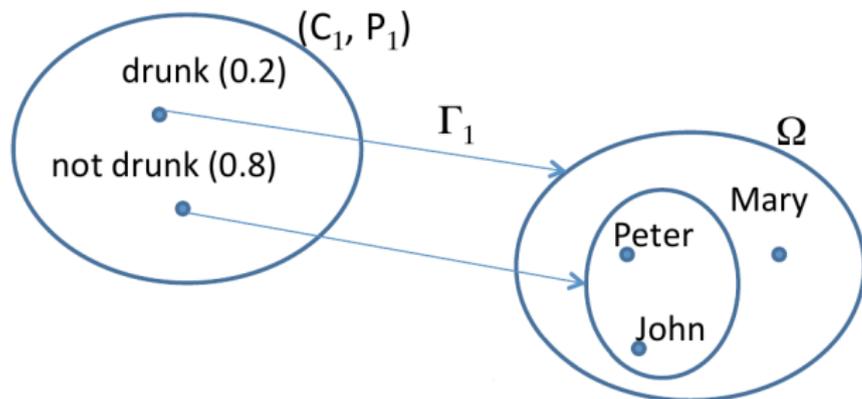
Example

- a mass function encodes evidence directly supporting **propositions** [Shafer, 1976] - let us see an example
- a murder has been committed. There are three suspects:
 $\Omega = \{Peter, John, Mary\}$
- available evidence: a witness saw the murderer going away, but he is short-sighted and he only saw that it was a man. We know that the witness is drunk 20 % of the time
- if the witness was not drunk, we know that $\omega \in \{Peter, John\}$
- otherwise, we only know that $\omega \in \Omega$. The first case holds with probability 0.8
- the corresponding mass function is:

$$m(\{Peter, John\}) = 0.8, \quad m(\Omega) = 0.2$$

Random set interpretation

of belief functions



- a mass assignment on Ω is induced by a probability distribution P on a different domain \mathcal{C} , via a multi-valued mapping $\Gamma : \mathcal{C} \rightarrow 2^\Omega$
- Γ maps *elements* $c \in \mathcal{C}$ ("codes") to *subsets* of Ω
- this is a **random set**, i.e., a set-valued random variable
- in the example, the source of the evidence is the probability P_1 that the witness is drunk (or not)
- Γ_1 maps $\{not\ drunk\} \in \mathcal{C}_1$ to $\{Peter, John\} \subset \Omega$

Belief (BFs) and plausibility functions

induced by a mass function

- for any $A \subseteq \Omega$, we can define:
 - the total **degree of support (belief)** in A as the probability that the evidence implies A :

$$Bel(A) = P(\{c \in \mathcal{C} \mid \Gamma(c) \subseteq A\}) = \sum_{B \subseteq A} m(B)$$

- the **plausibility** of A as the probability that the evidence does not contradict A :

$$Pl(A) = P(\{c \in \mathcal{C} \mid \Gamma(c) \cap A \neq \emptyset\}) = 1 - Bel(\bar{A})$$

- the uncertainty on the truth value of the proposition " $\omega \in A$ " is represented by two numbers: $Bel(A)$ and $Pl(A)$, with

$$Bel(A) \leq Pl(A)$$

Special cases of belief functions

they generalise both probabilities and possibilities

- if all focal sets of m are singletons, then m is said to be **Bayesian**
- it is equivalent to a **probability distribution**, and $Bel = Pl$ is a probability measure
- if the focal sets of m are nested, then m is said to be **consonant**
- in that case Pl is a **possibility measure**, i.e.,

$$Pl(A \cup B) = \max(Pl(A), Pl(B)), \quad \forall A, B \subseteq \Omega,$$

and Bel is the dual **necessity measure**

- the **contour function** $pl(\omega) = Pl(\{\omega\})$ corresponds to the possibility distribution (membership function)

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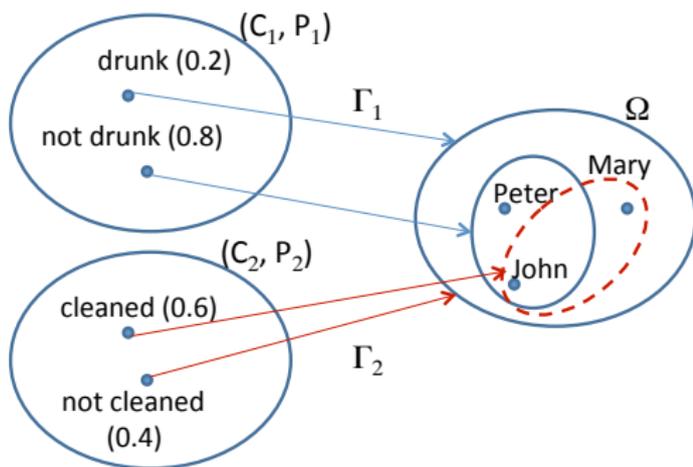
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Combination of evidence

Murder example continued

- when we have separate bodies of evidence, each represented by a belief function, **can we combine them** in order to estimate the state of the world, or make a decision?
- the first item of evidence gave us: $m_1(\{Peter, John\}) = 0.8$, $m_1(\Omega) = 0.2$
- new piece of evidence: a blond hair has been found
- also, there is a probability 0.6 that the room has been cleaned before the crime
- this second body of evidence is encoded by the mass assignment $m_2(\{John, Mary\}) = 0.6$, $m_2(\Omega) = 0.4$
- how to combine these two pieces of evidence?
- again, an answer can be given within the “random set” interpretation of belief functions

Combination of evidence



- if codes $c_1 \in C_1$ and $c_2 \in C_2$ were selected, $\omega \in \Gamma_1(c_1) \cap \Gamma_2(c_2)$
- if the codes are selected **independently**, then the probability that the pair (c_1, c_2) is selected is $P_1(\{c_1\})P_2(\{c_2\})$
- if $\Gamma_1(c_1) \cap \Gamma_2(c_2) = \emptyset$, (c_1, c_2) cannot be selected, hence:
- the joint probability distribution on $C_1 \times C_2$ must be conditioned to eliminate such pairs

Dempster's rule

Definition

- under these assumptions we get Dempster's rule of combination
- let m_1 and m_2 be two mass functions on the same frame Ω , induced by two **independent pieces of evidence**
- their combination using **Dempster's rule** is defined as:

$$(m_1 \oplus m_2)(A) = \frac{1}{1 - \kappa} \sum_{B \cap C = A} m_1(B)m_2(C), \quad \forall \emptyset \neq A \subseteq \Omega,$$

where

$$\kappa = \sum_{B \cap C = \emptyset} m_1(B)m_2(C)$$

is the **degree of conflict** between m_1 and m_2

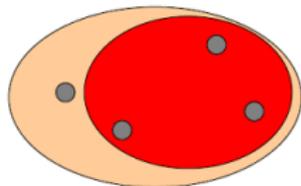
- their Dempster's sum $m_1 \oplus m_2$ exists iff $\kappa < 1$
- can be easily extended to any number of BFs

Dempster's rule

A simple numerical example

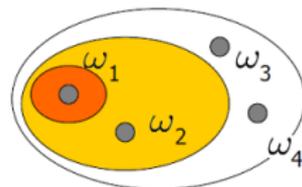
▪ $Bel_1:$

$$m(\{\omega_1\})=0.7, m(\{\omega_1, \omega_2\})=0.3$$



▪ $Bel_2:$

$$m(\Omega)=0.1, m(\{\omega_2, \omega_3, \omega_4\})=0.9$$

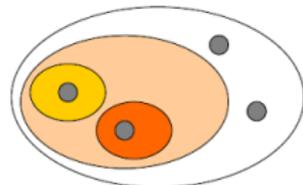


▪ $Bel_1 \oplus Bel_2:$

$$m(\{\omega_1\}) = 0.7 \cdot 0.1 / 0.37 = 0.19$$

$$m(\{\omega_2\}) = 0.3 \cdot 0.9 / 0.37 = 0.73$$

$$m(\{\omega_1, \omega_2\}) = 0.3 \cdot 0.1 / 0.37 = 0.08$$



Dempster's rule

Properties

- Dempster's rule has some interesting properties:
- commutativity, associativity, existence of a neutral element: the **vacuous** BF m_Ω
- it generalises set-theoretical **intersection**: if m_A and m_B are logical mass functions and $A \cap B \neq \emptyset$, then

$$m_A \oplus m_B = m_{A \cap B}$$

- it generalises **probabilistic conditioning** via Bayes' rule: if m is a Bayesian mass function and m_A is a logical mass function, then $m \oplus m_A$ is a Bayesian mass function that corresponding to Bayes' conditioning of m by A

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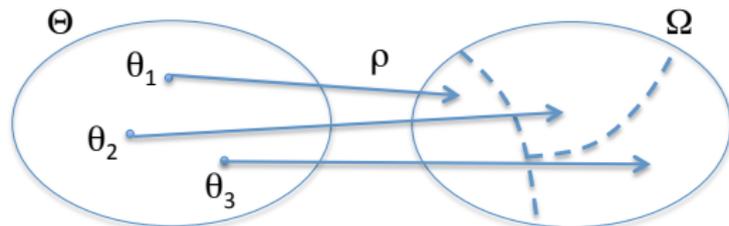
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Families of frames

Refinements and coarsenings

- the theory allows us to **handle evidence impacting on different but related domains**
- assume we are interested in the nature of an object in a road scene. We could describe it, e.g., in the frame $\Theta = \{\text{vehicle, pedestrian}\}$, or in the finer frame $\Omega = \{\text{car, bicycle, motorcycle, pedestrian}\}$
- other example: different image features in pose estimation
- a frame Ω is a **refinement** of a frame Θ (or, equivalently, Θ is a coarsening of Ω) if elements of Ω can be obtained by splitting some or all of the elements of Θ



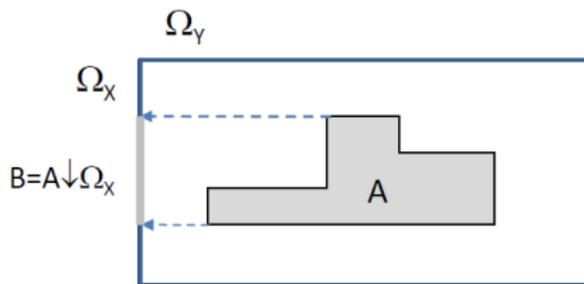
Compatible frames

Families of frames

- when Ω is a refinement for a collection $\Theta_1, \dots, \Theta_N$ of other frames it is called their **common refinement**
- two frames are said to be **compatible** if they do have a common refinement
- compatible frames can be associated with **different variables/attributes/features**:
 - let $\Omega_X = \{\text{red, blue, green}\}$ and $\Omega_Y = \{\text{small, medium, large}\}$ be the domains of attributes X and Y describing, respectively, the color and the size of an object
 - in such a case the common refinement $\Omega_X \times \Omega_Y$ is Ω_X and Ω_Y
- or, they can be descriptions of the **same variable at different levels of granularity** (as in the road scene example)
- **evidence can be moved from one frame to another** within a family of compatible frames

Marginalization

- let Ω_X and Ω_Y be two compatible frames
- let m^{XY} be a mass function on $\Omega_X \times \Omega_Y$
- it can be expressed in the coarser frame Ω_X by transferring each mass $m^{XY}(A)$ to the **projection** of A on Ω_X :



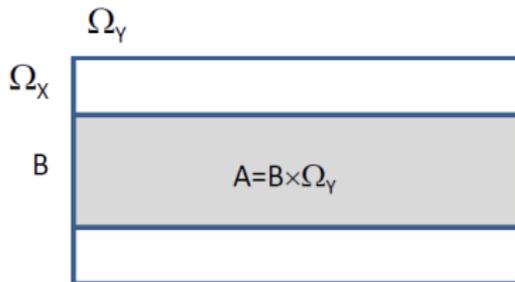
- we obtain a **marginal** mass function on Ω_X :

$$m^{XY \downarrow X}(B) = \sum_{\{A \subseteq \Omega_{XY}, A \downarrow \Omega_X = B\}} m^{XY}(A) \quad \forall B \subseteq \Omega_X$$

- (again, it generalizes both set projection and probabilistic marginalization)

Vacuous extension

- the “inverse” of marginalization
- a mass function m^X on Ω_X can be expressed in $\Omega_X \times \Omega_Y$ by transferring each mass $m_X(B)$ to the **cylindrical extension** of B :



- this operation is called the **vacuous extension** of m_X in $\Omega_X \times \Omega_Y$:

$$m^{X \uparrow XY}(A) = \begin{cases} m^X(B) & \text{if } A = B \times \Omega_Y \\ 0 & \text{otherwise} \end{cases}$$

- a strong feature of belief theory: the vacuous belief function (our representation of ignorance) **is left unchanged when moving from one compatible frame to another**

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The multiple semantics

of belief functions

- being complex objects, belief functions have a number of (sometimes conflicting) semantics and mathematical interpretations
- original one [Dempster 1967]: lower probabilities induced by a multivalued mapping
 - the mathematical representation: **random set** framework
- Shafer's (1976): representations of pieces of evidence in favour of propositions within someone's subjective state of belief
 - represented as **set functions** on a finite domain Ω
- as **convex sets of probability measures**, in a robust Bayesian interpretation
 - mathematically, a credal set whose lower and upper envelopes are belief and plausibility functions
- other equivalent mathematical formulations:
 - as non-additive (generalised) probabilities
 - as monotone capacities
 - as inner measures (linked to the rough set idea)

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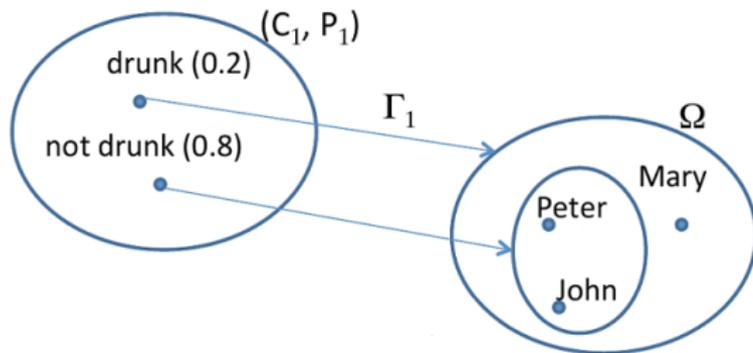
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Upper and lower probabilities

induced by multi-valued mappings

- Dempster has shown that mapping a probability distribution via a multi-valued map yields an object more general than a probability distribution: a belief function



- belief and plausibility values are interpreted as **lower and upper bounds** to the values of an unknown, underlying probability measure: $Bel(A) \leq P(A) \leq Pl(A)$ for all $A \subseteq \Omega$

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Belief functions as credal sets

or convex sets of probabilities

- each focal element A of mass $m(A)$ as the indication of the existence of a mass $m(A)$ “floating” inside A
- **constraint on the probability measure** on Ω : a distribution is “consistent” with Bel if it is obtained by redistributing the mass of each focal element to its singletons
- **set of probabilities consistent with b :**

$$\mathcal{P}[Bel] \doteq \left\{ P \in \mathcal{P} : P(A) \geq Bel(A) \forall A \subseteq \Omega \right\}$$

- it is a polytope in the probability simplex, with vertices induced by permutations of the elements of Ω
- Shafer disavowed any probability-bound interpretation
- also criticized by Walley as incompatible with Dempster’s rule of combination

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Belief functions as set functions

- Shafer's definition in terms of set functions [Aigner]
- a **belief function** $Bel : 2^\Omega \rightarrow [0, 1]$ is such that,

$$Bel(A) = \sum_{B \subseteq A} m(B)$$

where $m : 2^\Omega \rightarrow [0, 1]$ is a basic probability assignment s.t.

$$m(\emptyset) = 0, \sum_{A \subseteq \Omega} m(A) = 1, m(A) \geq 0 \forall A \subseteq \Omega$$

- operating with belief functions reduces to manipulating their focal elements
- in Shafer's framework, the mass assignment is derived by "impact of evidence" associated with propositions via an exponential-like relation

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As non-additive probabilities

or “generalised” probabilities

Probability measure

A function $P : \mathbf{F} \rightarrow [0, 1]$ over a σ -field $\mathbf{F} \subset 2^\Omega$ such that

- $P(\emptyset) = 0, P(\Omega) = 1$;
- if $A \cap B = \emptyset, A, B \in \mathbf{F}$ then $P(A \cup B) = P(A) + P(B)$ (*additivity*).
- if we relax the third constraint to allow the function to meet additivity **only as a lower bound** we obtain a:

Belief function

A function $Bel : 2^\Omega \rightarrow [0, 1]$ from the power set 2^Ω to $[0, 1]$ such that:

- $Bel(\emptyset) = 0, Bel(\Omega) = 1$;
- for every n and for every collection $A_1, \dots, A_n \in 2^\Omega$ we have that:

$$Bel(A_1 \cup \dots \cup A_n) \geq \sum_i Bel(A_i) - \sum_{i < j} Bel(A_i \cap A_j) + \dots \\ \dots + (-1)^{n+1} Bel(A_1 \cap \dots \cap A_n)$$

Belief functions as completely monotone capacities

- a function $Bel : 2^\Omega \rightarrow [0, 1]$ is a **completely monotone capacity**, i.e., it verifies $Bel(\emptyset) = 0$, $Bel(\Omega) = 1$ and

$$Bel\left(\bigcup_{i=1}^k A_i\right) \geq \sum_{\emptyset \neq I \subseteq \{1, \dots, k\}} (-1)^{|I|+1} Bel\left(\bigcap_{i \in I} A_i\right)$$

for any $k \geq 2$ and for any family A_1, \dots, A_k in 2^Ω

- conversely, to any completely monotone capacity Bel corresponds a **unique mass function** m such that:

$$m(A) = \sum_{\emptyset \neq B \subseteq A} (-1)^{|A|-|B|} Bel(B), \quad \forall A \subseteq \Omega$$

- in combinatorics this is called **Moebius transform**
- m , Bel and Pl are thus **equivalent representations** of the same piece of evidence

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Belief functions as random sets

Rationale

- given a multi-valued mapping Γ , a straightforward step is to consider the probability value $P(c)$ as attached to the subset $\Gamma(c) \subseteq \Omega$
- what we obtain is a **random set** in Ω , i.e., a probability measure on a collection of subsets
- roughly speaking, a random set is a set-valued random variable
- the degree of belief $Bel(A)$ of an event A becomes the **cumulative distribution function (CDF) of the open interval of sets** $\{B \subseteq A\}$
- this approach has been emphasized in particular by [Nguyen,1978] and [Hestir,1991] and [Shafer,1987]
- example: **a dice where one or more of faces are covered** so that we do not know what's beneath is a random variable which "spits" subsets of possible outcomes: a random set

Belief functions as random sets

Mathematics

- the **lower inverse** of Γ is defined as:

$$\Gamma_*(A) \doteq \{c \in \mathcal{C} : \Gamma(c) \subset A, \Gamma(c) \neq \emptyset\}$$

while its **upper inverse** is

$$\Gamma^*(A) \doteq \{c \in \mathcal{C} : \Gamma(c) \cap A \neq \emptyset\}$$

- given two σ -fields \mathcal{A}, \mathcal{B} on \mathcal{C}, Ω respectively, Γ is said **strongly measurable** iff $\forall B \in \mathcal{B}, \Gamma^*(B) \in \mathcal{A}$
- the **lower probability measure** on \mathcal{B} is defined as $P_*(B) \doteq P(\Gamma_*(B))$ for all $B \in \mathcal{B}$ - this is nothing but a belief function!
- Nguyen proved that, if Γ is strongly measurable, the CDF \hat{P} of the random set coincides with the lower probability measure:

$$\hat{P}[I(B)] = P_*(B) \quad \forall B \in \mathcal{B}, \quad I(B) \doteq \{C \in \mathcal{B}, C \subseteq B\}$$

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The inference problem

Building belief functions from the available data

- first step in any estimation/decision problem: **constructing a belief function from the available evidence**
- belief functions can be constructed from both statistical data (quantitative inference) and experts' preferences (qualitative inference)
- inference from **statistical data**: we will see two
 - Dempster's approach
 - likelihood-based approach
- inference from **qualitative data**
 - Wong and Lingras's perceptron idea
 - Qualitative Discrimination Process (QDP)
 - Ben Yaghlane's constrained optimisation

Inferring belief functions from statistical data

- consider a **statistical model**

$$\{f(x; \theta), x \in \mathbb{X}, \theta \in \Theta\},$$

where \mathbb{X} is the sample space and Θ the parameter space

- having observed x , how to **quantify the uncertainty about the parameter** θ , without specifying a prior probability distribution?
- two main approaches using belief functions:
 - Dempster's approach** based on an auxiliary variable with a pivotal probability distribution [Dempster, 1967]
 - Likelihood-based approach** [Shafer, 1976, Wasserman 1990]

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Sampling model

Dempster's approach to statistical inference

- suppose that the sampling model $X \sim f(x; \theta)$ can be represented by an “a-equation” of the form

$$X = a(\theta, U)$$

where $U \in \mathbb{U}$ is an **(unobserved) auxiliary variable** with known probability distribution μ independent of θ

- this representation is quite natural in the context of sampling and **data generation**
- for instance, to generate a continuous random variable X with cumulative distribution function (CDF) F_θ , one might draw U from $\mathcal{U}([0, 1])$ and set

$$X = F_\theta^{-1}(U)$$

From a -equations to belief functions

- the equation $X = a(\theta, U)$ defines a multi-valued mapping (a “compatibility relation”)

$$\Gamma : U \rightarrow \Gamma(U) = \left\{ (X, \theta) \in \mathbb{X} \times \Theta \mid X = a(\theta, U) \right\}$$

- under the usual measurability conditions, the probability space $(\mathbb{U}, \mathcal{B}(\mathbb{U}), \mu)$ and the multi-valued mapping Γ induce a belief function $Bel_{\Theta \times \mathbb{X}}$ on $\mathbb{X} \times \Theta$
- conditioning (by Dempster's rule) $Bel_{\Theta \times \mathbb{X}}$ on θ yields the desired sampling distribution $f(\cdot; \theta)$ on \mathbb{X}
- conditioning it on $X = x$ gives a belief function $Bel_{\Theta}(\cdot; x)$ on Θ

Example: Bernoulli sample

Dempster's approach to inference

- let $X = (X_1, \dots, X_n)$ consist of **independent Bernoulli observations** and $\theta \in \Theta = [0, 1]$ is the probability of success

- sampling model:

$$X_i = \begin{cases} 1 & \text{if } U_i \leq \theta \\ 0 & \text{otherwise,} \end{cases}$$

where $U = (U_1, \dots, U_n)$ has pivotal measure $\mu = \mathcal{U}([0, 1]^n)$

- having observed the number of successes $y = \sum_{i=1}^n x_i$, the belief function $Bel_{\Theta}(\cdot; x)$ is induced by a **random closed interval**

$$[U_{(y)}, U_{(y+1)}],$$

where $U_{(i)}$ denotes the i -th order statistics from U_1, \dots, U_n

- quantities like $Bel_{\Theta}([a, b]; x)$ or $Pl_{\Theta}([a, b]; x)$ are readily calculated

Discussion

Dempster's method

- Dempster's model has several nice features:
 - it allows us to quantify the uncertainty on Θ after observing the data, without having to specify a prior distribution on Θ
 - when a Bayesian prior P_0 is available, **combining it with $Bel_{\Theta}(\cdot; x)$ using Dempster's rule yields the Bayesian posterior:**

$$Bel_{\Theta}(\cdot; x) \oplus P_0 = P(\cdot|x)$$

- it also has some drawbacks:
 - it often leads to **cumbersome or even intractable calculations** except for very simple models, which imposes the use of Monte-Carlo simulations (see Computation later)
 - more fundamentally, **the analysis depends on the a-equation $X = a(\theta, U)$ and the auxiliary variable U** , which are not unique for a given statistical model $\{f(\cdot; \theta), \theta \in \Theta\}$
 - As U is not observed, how can we argue for an a-equation or another?

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Likelihood-based belief function

Requirements

- 1 **Likelihood principle:** $Bel_{\Theta}(\cdot; x)$ should be based only on the likelihood function $L(\theta; x) = f(x; \theta)$
- 2 **Compatibility with Bayesian inference:** when a Bayesian prior P_0 is available, combining it with $Bel_{\Theta}(\cdot, x)$ using Dempster's rule should yield the Bayesian posterior:

$$Bel_{\Theta}(\cdot; x) \oplus P_0 = P(\cdot|x)$$

- 3 **Principle of minimal commitment:** among all the belief functions satisfying the previous two requirements, $Bel_{\Theta}(\cdot; x)$ should be the least committed (least informative)

Likelihood-based belief function

Solution

- $Bel_{\Theta}(\cdot; x)$ is the **consonant belief function** with contour function (plausibility of singletons) equal to the **normalized likelihood**:

$$pl(\theta; x) = \frac{L(\theta; x)}{\sup_{\theta' \in \Theta} L(\theta'; x)}$$

- the corresponding plausibility function is:

$$Pl_{\Theta}(A; x) = \sup_{\theta \in A} pl(\theta; x) = \frac{\sup_{\theta \in A} L(\theta; x)}{\sup_{\theta \in \Theta} L(\theta; x)}, \quad \forall A \subseteq \Theta$$

- the corresponding random set is: $(\Omega, \mathcal{B}(\Omega), \mu, \Gamma_x)$ with $\Omega = [0, 1]$, $\mu = \mathcal{U}([0, 1])$ and

$$\Gamma_x(\omega) = \left\{ \theta \in \Theta \mid pl(\theta; x) \geq \omega \right\}$$

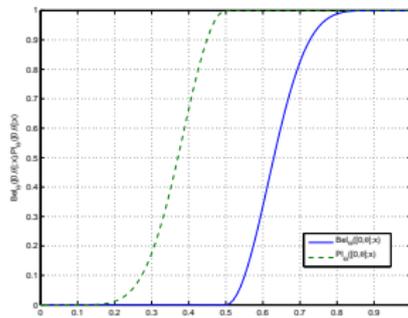
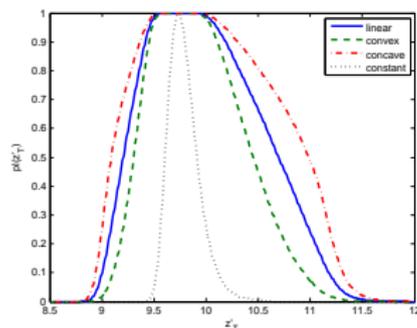
Example: Bernoulli sample

- let $X = (X_1, \dots, X_n)$ consist of independent Bernoulli observations and $\theta \in \Theta = [0, 1]$ is the probability of success
- we get

$$pl(\theta; \mathbf{x}) = \frac{\theta^y (1 - \theta)^{n-y}}{\hat{\theta}^y (1 - \hat{\theta})^{n-y}},$$

where $y = \sum_{i=1}^n x_i$ and $\hat{\theta}$ is the MLE

- example for $n = 20$ and $y = 10$:



Discussion

Likelihood method

- the likelihood-based method is much simpler to implement than Dempster's method, even for complex models.
- by construction, it **boils down to Bayesian inference when a Bayesian prior is available**
- it is compatible with usual likelihood-based inference:
 - assume that $\theta = (\theta_1, \theta_2) \in \Theta_1 \times \Theta_2$ and θ_2 is a **nuisance parameter**. The marginal contour function on Θ_1

$$pl(\theta_1; x) = \sup_{\theta_2 \in \Theta_2} pl(\theta_1, \theta_2; x) = \frac{\sup_{\theta_2 \in \Theta_2} L(\theta_1, \theta_2; x)}{\sup_{(\theta_1, \theta_2) \in \Theta} L(\theta_1, \theta_2; x)}$$

is the relative **profile likelihood** function

- the plausibility of a composite hypothesis $H_0 \subset \Theta$

$$Pl(H_0; x) = \frac{\sup_{\theta \in H_0} L(\theta; x)}{\sup_{\theta \in \Theta} L(\theta; x)}$$

is the usual **likelihood ratio statistics** $\Lambda(x)$

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Wong and Lingras

Building belief functions from preferences

- Wong and Lingras [16] proposed a method for generating BFs from a body of qualitative preference relations between propositions
- two binary relations: preference $\cdot >$ and indifference \sim
- goal: to build a belief function Bel such that $A > B$ iff $Bel(A) > Bel(B)$ and $A \sim B$ iff $Bel(A) = Bel(B)$
- exists if $\cdot >$ is a weak order and \sim an equivalence relation
- **Algorithm**
 - 1 consider all propositions that appear in the preference relations as potential focal elements (FEs)
 - 2 elimination: if $A \sim B$ for some $B \subset A$ then A is not a FE
 - 3 a perceptron algorithm is used to generate the mass m by solving the system of remaining equalities and disequalities
- however: it selects arbitrarily one solution over many
- does not address possible inconsistency in the given preferences

Ben Yaghlane's constrained optimisation approach

Building belief functions from preferences

- uses preferences and indifferences as in Wong and Lingras, with same axioms..
- .. but converts them into a **constrained optimisation problem**
- objective function: maximise the entropy/uncertainty of the BF to generate (least informative result)
- constraints derived from input preferences/indifferences, i.e.

$$A \succ B \leftrightarrow Bel(A) - Bel(B) \geq \epsilon, \quad A \sim B \leftrightarrow |Bel(A) - Bel(B)| \leq \epsilon$$

- ϵ is a constant specified by the expert
- various uncertainty measures can be plugged in (see Advances)

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Conditional belief functions

A variety of proposals

- many different approaches to conditioning belief functions have been proposed
- a non-exhaustive list:
 - original Dempster's conditioning
 - **lower and upper envelopes** of conditional probabilities [Fagin and Halpern]
 - **geometric conditioning** [Suppes]
 - **unnormalized** conditional belief functions [Smets]
 - generalised Jeffrey's rules [Smets]
 - **sets of equivalent events** under multi-valued mappings [Spies]
 - conditioning by distance minimisation [Cuzzolin]
- implications of the notion of conditional belief function:
 - **generalised Bayes theorem** [Smets]
 - the generalisation of the **total probability** theorem

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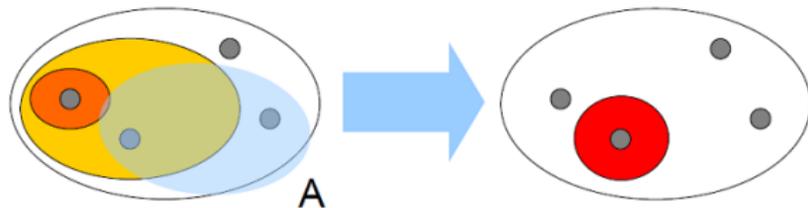
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Dempster's conditioning

Conditioning

- Dempster's rule of combination is associated with a conditioning operator
- suppose we have an "a-priori" BF Bel
- given a new event A , the "logical" belief function such that $m(A) = 1$ can be defined ...
- ... and combined with Bel using Dempster's rule
- the resulting BF is the conditional belief function given A , *a la Dempster*



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Lower conditional envelopes

Conditioning

- Fagin and Halpern proposed an approach based on interpretation of a belief function as the lower envelope of the family of probabilities consistent with it (robust Bayesian)

$$Bel(A) = \inf_{P \in \mathcal{P}[Bel]} P(A)$$

- they define the conditional belief as the **lower envelope** (that is, the infimum) **of the family of conditional probability functions** $P(A|B)$, where P is consistent with Bel :

$$Bel(A|B) \doteq \inf_{P \in \mathcal{P}[Bel]} P(A|B), \quad Pl(A|B) \doteq \sup_{P \in \mathcal{P}[Bel]} P(A|B)$$

- trivially generalises conditional probability
- have been considered by other authors too, e.g. Dempster 1967 and Walley 1981

Lower conditional envelopes

versus Dempster's

- the authors provide a closed-form expression for it:

$$Bel(A|B) = \frac{Bel(A \cap B)}{Bel(A \cap B) + Pl(A \cap B)}, \quad Pl(A|B) = \frac{Pl(A \cap B)}{Pl(A \cap B) + Bel(A \cap B)}$$

- lower/upper envelopes of arbitrary sets of probabilities are not in general belief functions, but **these actually are**, as Fagin and Halpern have proven
- they are quite different from Dempster's conditioning:

$$Bel_{\oplus}(A|B) = \frac{Bel(A \cup \bar{B})}{1 - Bel(\bar{B})}, \quad Pl_{\oplus}(A|B) = \frac{Pl(A \cap B)}{Pl(B)}$$

- in fact, they provide a **more conservative estimate**:

$$Bel(A|B) \leq Bel_{\oplus}(A|B) \leq Pl_{\oplus}(A|B) \leq Pl(A|B)$$

- Fagin and Halpern argue that Dempster's conditioning behaves unreasonably on their "three prisoners" example

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Revision versus focussing

in belief as opposed to probability theory

Focussing

No new information is introduced, we merely focus on a specific subset of the original set.

Belief revision

A state of belief is modified to take into account a new piece of information.

- in probability theory, **both are expressed by Bayes' rule**, but they are conceptually different operations
- in belief theory, these principles **lead to different conditioning rules**
- the application of revision and focussing to belief theory has been explored by Smets in his Transferable Belief Model (TBM)
- here we are not assuming any random set generating *Bel*, nor any underlying convex sets of probabilities

Suppes' geometric conditioning

Conditioning

- the **geometric conditioning** proposed by Suppes and Zanotti

$$Bel_G(A|B) = \frac{Bel(A \cap B)}{Bel(B)},$$

is indeed a consequence of the **focussing** idea

- (this was proved by Smets using the “probability of provability” interpretation of belief functions, yes, yet another one!)
- somewhat dual to Dempster's conditioning, as it **replaces probability with belief** in Bayes' rule
- remember that Dempster's rule dually **replaces probability with plausibility** in Bayes' rule

$$Pl_{\oplus}(A|B) = \frac{Pl(A \cap B)}{Pl(B)} \leftrightarrow Bel_G(A|B) = \frac{Bel(A \cap B)}{Bel(B)}$$

Smets' unnormalised rule

of conditioning

- to rebuke Bayesian criticisms, in his TBM Smets rejects the existence of a probability measure on a parent space \mathcal{C}
- Smets' (Dempster's) **unnormalized conditional belief function**:

$$m_U(\cdot|B) = \begin{cases} \sum_{X \subseteq B^c} m(A \cup X) & \text{if } A \subseteq B \\ 0 & \text{elsewhere} \end{cases}$$

- (in the TBM BFs which assign mass to \emptyset can exist, under the “open world” assumption)
- in terms of plausibilities: $Pl_U(A|B) = Pl(A \cap B)$ - in the TBM the mass $m(A)$ is transferred by conditioning on B to $A \cap B$
- it is a consequence of **belief revision** principles [Gilboa, Perea]

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Spies' sets of equivalent events under multi-valued mappings

Conditioning

- intriguing approach to conditioning, in the random set interpretation with $(\mathcal{C}, \mathcal{F}, P)$ and $\Gamma : \mathcal{C} \rightarrow 2^\Omega$
- null sets for $P(\cdot|A)$: $\mathcal{N}(P(\cdot|A)) = \{B \in \mathcal{A} : P(B|A) = 0\}$
- let Δ be the symmetric difference $A\Delta B = (A \cap \bar{B}) \cup (\bar{A} \cap B)$
- two events have the same conditional probability if they both are the symmetric difference between a same event and some null set
- a **conditional event** $[B|A]$ with $A, B \subseteq \Omega$ is a set of events with the same conditional probability $P(B|A)$:

$$[B|A] = B\Delta\mathcal{N}(P_A)$$

- you can prove that $[B|A] = \{C : A \cap B \subseteq C \subseteq \bar{A} \cup B\}$

Conditional belief functions

Spies' approach

- by applying to conditional events a multivalued mapping Spies gave a new definition of conditional belief function
- **conditional multivalued mapping** for $B \subseteq \Omega$: $\Gamma_B(c) = [\Gamma(c)|B]$, where $\Gamma : \mathcal{C} \rightarrow 2^\Omega$
- (if $A = \Gamma(c)$, Γ_B maps c to $[A|B]$)
- consequence: to all elements of each conditioning event (an equivalence class) must be assigned equal belief/plausibility
- a **conditional belief function** is then a “second-order” BF with values on *collections* of focal elements (the conditional events)

$$\text{Bel}([C|B]) = P(\{c : \Gamma_B(c) = [C|B]\}) = \frac{1}{K} \sum_{A \in [C|B]} m(A)$$

- it is *not* a BF on the sub-algebra $\{Y = C \cap B, C \subseteq \Omega\}$
- Spies' conditional belief functions are **closed under Dempster's rule** of combination

Jeffrey's rule of conditioning

or Total Probability Theorem

- suppose P is defined on a σ -algebra \mathbb{A}
- there is a new prob measure P' on a sub-algebra \mathbb{B} of \mathbb{A} , and the updated probability P'' has to:

- 1 meet the prob values specified by P' for events in \mathbb{B}
- 2 be such that $\forall B \in \mathbb{B}, X, Y \subset B, X, Y \in \mathbb{A}$

$$\frac{P''(X)}{P''(Y)} = \begin{cases} \frac{P(X)}{P(Y)} & \text{if } P(Y) > 0 \\ 0 & \text{if } P(Y) = 0 \end{cases}$$

- there is a unique solution:

$$P''(A) = \sum_{B \in \mathbb{B}} P(A|B)P'(B)$$

- meaning: the initial probability **stands corrected by the second one** on a number of events
- generalises conditioning (obtained when $P'(B) = 1$ for some B)

Jeffrey's rule generalised

Jeffrey's rule for belief functions

- 1 Let $\Pi = \{B_1, \dots, B_n\}$ a disjoint partition of Ω ;
- 2 m_1, \dots, m_n the BPAs of BF's conditional on B_1, \dots, B_n respectively;
- 3 m_B an unconditional belief function on the coarsening associated with the partition Π

Then the belief function $Bel_{tot}(A) = \sum_{C \subseteq A} (m_B \oplus \oplus_i^n m_{B_i})(C)$ is a marginal belief function on Ω , and if all BF's are probabilities is reduces to the result of Jeffrey's rule of total probability

- combining the a-priori with all the conditionals we get a marginal
- is this the only solution? we will discuss this later (total belief theorem)

Conditioning approaches

A summary

- various approaches to conditioning have been proposed:

- Dempster's** (normalised) conditioning:

$$Bel_{\oplus}(A|B) = \frac{Bel(A \cup \bar{B})}{1 - Bel(\bar{B})}, \quad Pl_{\oplus}(A|B) = \frac{Pl(A \cap B)}{Pl(B)}$$

- lower and upper **conditional envelopes**:

$$Bel(A|B) = \frac{Bel(A \cap B)}{Bel(A \cap B) + Pl(\bar{A} \cap B)}, \quad Pl(A|B) = \frac{Pl(A \cap B)}{Pl(A \cap B) + Bel(\bar{A} \cap B)}$$

- geometric** conditioning:

$$Bel_G(A|B) = \frac{Bel(A \cap B)}{Bel(B)},$$

- Smets' **unnormalised** rule: $Pl_U(A|B) = Pl(A \cap B)$

- Spies' conditioning: $Bel[A|B] \propto \sum_{X: A \cap B \subseteq X \subseteq A \cup \bar{B}} m(X)$

- derived from **different revision principles**
 - follow **different semantic interpretations** (TBM rather than random set, open versus closed world assumption, robust Bayesian)

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Bayes' theorem

generalised to belief functions

- have a conditional probability $P(x|\theta_i)$ over observations $x \in X$, and an a-priori probability P_0 over a set of hidden variables $\theta_i \in \Theta$
- (for instance, x is a symptom and θ_i a disease)
- after observing x , the probability distribution on Θ is updated to the posterior via Bayes's theorem:

$$P(\theta_i|x) = \frac{P(x|\theta_i)P_0(\theta_i)}{\sum_j P(x|\theta_j)P_0(\theta_j)} \quad \forall \theta_j \in \Theta$$

- the **GBT is a generalisation of Bayes' theorem for conditional BFs, when the a-priori BF on Θ is vacuous**
- Dempster's normalised/unnormalised conditioning is assumed
- (a further generalisation for non-vacuous priors is proposed in Smets' work) [Smets 1993]

Cognitive independence

- consider a belief function over the product space $X \times Y$
- the two variables are **cognitively independent** if

$$pl_{X \times Y}(x \cap y) = pl_X(x)pl_Y(y) \quad \forall x \subseteq X, y \subseteq Y$$

- cognitive independence extends stochastic independence
- **conditional** cognitive independence reads as

$$pl_{X \times Y}(x \cap y | \theta_i) = pl_X(x | \theta_i)pl_Y(y | \theta_i) \quad \forall x, y, \theta_i$$

and implies that the ratio of plausibility/belief on X does not depend on Y :

$$\frac{pl_X(x_1 | y)}{pl_X(x_2 | y)} = \frac{pl_X(x_1)}{pl_X(x_2)}, \quad \frac{Bel_X(x_1 | y)}{Bel_X(x_2 | y)} = \frac{Bel_X(x_1)}{Bel_X(x_2)}$$

Likelihood principle

Edwards, 1929

- the likelihood of an hypothesis given the data amounts to the conditional probability of the data given the hypothesis:

$$l(\theta_i|x) = p(x|\theta_i)$$

and, for unions of singleton hypotheses:

$$l(\theta = \{\theta_1, \dots, \theta_k\}|x) = \max \left\{ l(\theta_i|x) : \theta_i \in \theta \right\}$$

- Shafer's somewhat similar proposal for statistical inference (see inference-likelihood method):

$$pl(\theta|x) = \max_{\theta_i \in \theta} pl(\theta_i|x)'$$

was rejected by Smets, for not satisfying the condition that, if two pieces of evidence are conditionally independent, $Bel_{\Theta}(\cdot|x, y)$ is the conjunctive combination of $Bel_{\Theta}(\cdot|x)$ and $Bel_{\Theta}(\cdot|y)$

Generalised Likelihood Principle

Smets' Generalised Likelihood Principle (GLP)

- 1 $pl_{\Theta}(\theta|x) = pl_x(x|\theta)$
- 2 For all x, θ the plausibility of data $pl(x|\theta)$ given a compound hypothesis $\theta = \{\theta_1, \dots, \theta_m\}$ is a function of only

$$\{pl(x|\theta_i), pl(\bar{x}|\theta_i) : \theta_i \in \theta\}$$

- the form of the function is not assumed (not necessarily the max)
- both $pl(x|\theta_i)$ and $pl(\bar{x}|\theta_i)$ are necessary because of the non-additivity of belief functions
- justified by the following requirements:
 - $pl(x|\theta)$ remains the same on the coarsening of X formed by just x and \bar{x}
 - plausibilities for $\theta_j \notin \theta$ are irrelevant for $pl(x|\theta)$

Generalised Bayesian Theorem

and the disjunctive rule of combination

- under conditional cognitive independence and the Generalised Likelihood Principle (2), $Bel_X(\cdot|\theta)$, $\theta \subset \Theta$ is generated from the $\{Bel_X(\cdot|\theta_i), \theta_i \in \Theta\}$ by **disjunctive rule of combination**

$$Pl_X(x|\theta) = 1 - \prod_{\theta_i \in \theta} (1 - Pl_X(x|\theta_i)), \quad Bel_X(x|\theta) = \prod_{\theta_i \in \theta} Bel_X(x|\theta_i)$$

- then, condition (1) of the GLP $p_{l\Theta}(\theta|x) = pl_X(x|\theta)$ implies the **generalised Bayes theorem**:

$$Pl_{\Theta}(\theta|x) = \frac{1}{K} \left(1 - \prod_{\theta_i \in \theta} (1 - pl_X(x|\theta_i)) \right)$$

$$Bel_{\Theta}(\theta|x) = \frac{1}{K} \left(\prod_{\theta_i \in \bar{\theta}} Bel_X(\bar{x}|\theta_i) - \prod_{\theta_j \in \theta} Bel_X(\bar{x}|\theta_j) \right)$$

where $K = 1 - \prod_{\theta_i \in \Theta} (1 - pl_X(x|\theta_i))$

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The total belief theorem

Generalising total probability to belief functions

Theorem

Suppose Θ and Ω are two frames of discernment, and $\rho : 2^\Omega \rightarrow 2^\Theta$ the unique refining between them. Let Bel_0 be a belief function defined over $\Omega = \{\omega_1, \dots, \omega_{|\Omega|}\}$. Suppose there exists a collection of belief functions $Bel_i : 2^{\Pi_i} \rightarrow [0, 1]$, where $\Pi = \{\Pi_1, \dots, \Pi_{|\Omega|}\}$, $\Pi_i = \rho(\{\omega_i\})$, is the partition of Θ induced by its coarsening Ω .

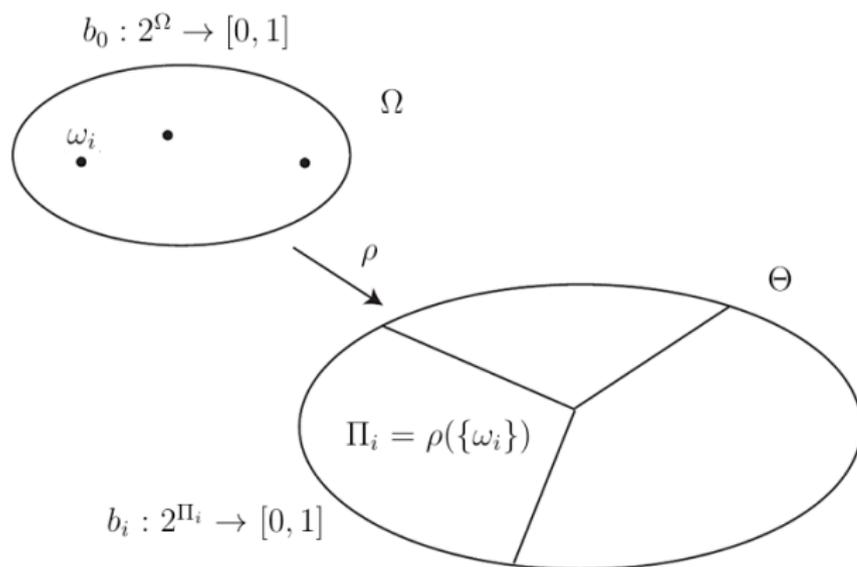
Then, there exists a belief function $Bel : 2^\Theta \rightarrow [0, 1]$ such that:

- 1 Bel_0 is the restriction of Bel to Ω
- 2 $Bel \oplus Bel_{\Pi_i} = Bel_i \forall i = 1, \dots, |\Omega|$, where Bel_{Π_i} is the logical belief function with mass

$$m_{\Pi_i}(A) = 1 \text{ } A = \Pi_i, \text{ } 0 \text{ otherwise}$$

The total belief theorem

Visual representation

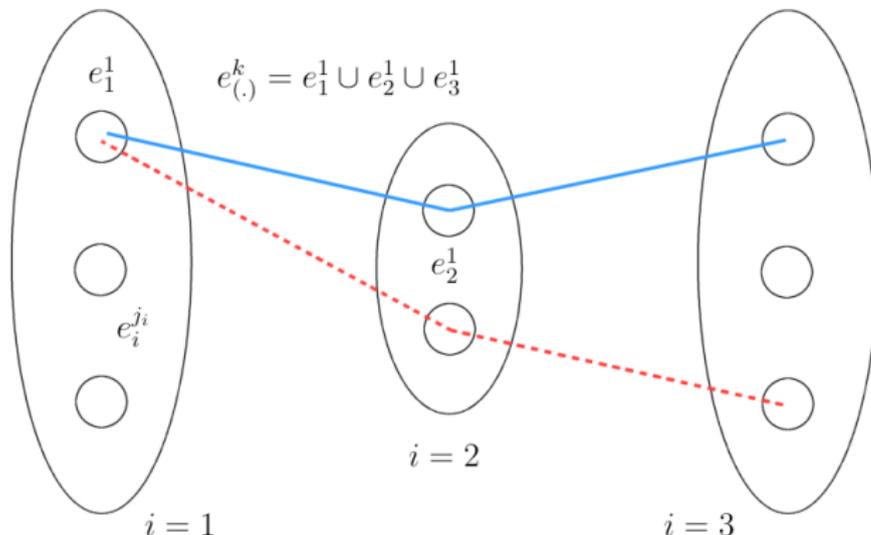


- pictorial representation of the total belief theorem

Structure of the focal elements

of the total belief function

- **restricted total belief theorem:** Bel_0 has only disjoint FEs
- pictorial representation of the structure of the FEs of a total BF Bel lying in the image of a focal element of Bel_0 of cardinality 3



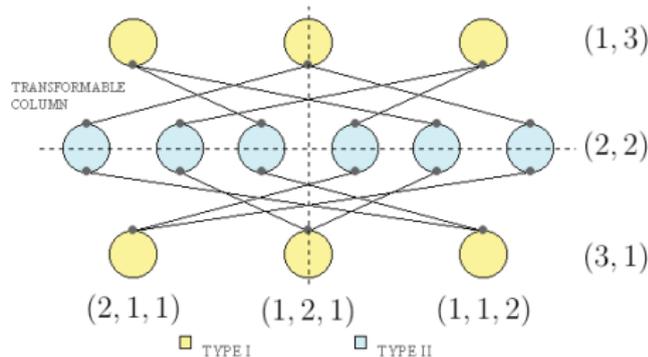
Graph of solutions

- potential solutions correspond to square linear systems, and form a graph whose nodes are linked by linear transformations of columns

$$e \mapsto e' = -e + \sum_{i \in \mathcal{C}} e_i - \sum_{j \in \mathcal{S}} e_j$$

where \mathcal{C} is a covering set for e (i.e., every component of e is covered by at least one of them), \mathcal{S} a set of selection columns

- at each transformation, the most negative component decreases



- general solution based on simplex-like optimisation?

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Efficient computation

with belief functions

- as belief functions are set functions, their complexity is exponential on the size n of the domain they are defined on
- combining belief functions via Dempster's rule is also exponential: $(2^n)^N$, where N is the number of BFs involved
- efficient approaches based on approximating the original evidence - in particular, approaches that **transform a belief function into a less complex uncertainty measure**
 - probability (Bayesian) transformation
 - possibility (consonant) transformation
- approaches based on the **local propagation** of evidence
 - Barnett's approximation
 - hierarchical evidence
 - Cano's propagation on DAGs
 - Shafer-Shenoy architecture
- **Monte-Carlo methods** [Wilson and Moral]

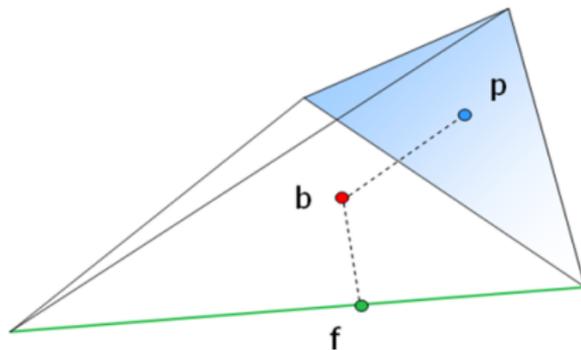
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Probability transformations

Mapping belief functions to probabilities

- **probability transform** of belief functions: an operator $pt : \mathcal{B} \rightarrow \mathcal{P}$, $b \mapsto pt[b]$ mapping belief measures onto probability distributions
- (not necessarily an element of the corresponding credal set)
- a number of transforms proposed, either as efficient implementations of ToE or tools for decision making
 - **pignistic** transform, central in the TBM [Smets]
 - **plausibility** and belief transform [Voorbraak, Cobb & Shenoy]
 - **orthogonal projection** and **intersection probability** [Cuzzolin]



Transformation in the TBM: the pignistic transform

Probability transformation

- Smets' Transferable Belief Model -> decisions made via "pignistic transform" ..
- .. resulting in a **pignistic probability**:

$$\text{Bet}P[b](x) = \sum_{A \ni x} \frac{m_b(A)}{|A|},$$

- its purpose is to allow decision making at the level of probabilities, typically in an expected utility framework
- the result of a redistribution process in which the mass of each focal element A is re-assigned to all its elements $x \in A$ on an equal basis
- it commutes with affine combination and is the **center of mass of the credal set of consistent probabilities**

Plausibility and belief transform

Probability transformation

- **plausibility transform** [Voorbraak 89]: maps a belief function to the **relative belief of singletons**:

$$\tilde{pl}(x) = \frac{pl(x)}{\sum_{y \in \Theta} pl(y)}$$

- relative plausibility of singletons \tilde{pl} is a perfect representative of Bel when combined with other probabilities by Dempster's rule \oplus
- meets a number of properties w.r.t. \oplus [Coob&Shenoy 03]
- the **relative belief transform** maps each belief function to the corresponding *relative belief of singletons*:

$$\tilde{bel}(x) = \frac{Bel(\{x\})}{\sum_{y \in \Theta} Bel(\{y\})}$$

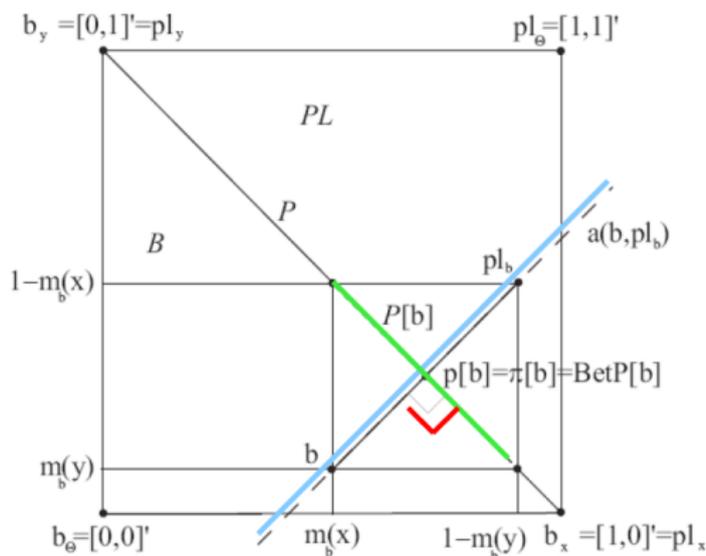
- first proposed by Daniel in 2006, its geometry and that of plausibility transform analyzed in [Cuzzolin 2010 AMAI]

Geometric transformations

Probability transformation

- the probability transformation problem can be posed in geometric terms [IEEE SMC-B07]

- orthogonal projection** $\pi[b] \rightarrow$ minimises the L2 distance from \mathcal{P}
- intersection probability** $p[b]$ \rightarrow intersection with \mathcal{P} of the belief-plausibility line



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the working
scientist

F. Cuzzolin
and
T. Denoëux

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Possibility transformations

Outer approximations

- necessity measures have as counterparts in the ToE **consonant** belief functions, whose focal elements are nested: $A_1 \subset \dots \subset A_m$, $A_i \subseteq \Theta$
- outer consonant approximations** [Dubois&Prade 90]: consonant BFs co which are dominated by the original BF on all events:

$$co(A) \leq Bel(A) \quad \forall A \subseteq \Theta$$

- for each possible maximal chain $A_1 \subset \dots \subset A_n$, $|A_i| = i$ of focal elements the **maximal outer consonant** approximation has mass

$$m_{\max}(A_i) = Bel(A_i) - Bel(A_{i-1})$$

- mirrors the behavior of the vertices of the credal set of probabilities dominating a belief function [Chateauneuf, Miranda& Grabish]

Possibility transformation

Isopignistic approximation

- completely different approximation in Smets' Transferable Belief Model [Smets94,05]
- **isopignistic" approximation:** the unique consonant belief function whose pignistic probability $BetP$ is identical to that of Bel [Dubois, Aregui]
- its contour function is:

$$pl_{iso}(x) = \sum_{x' \in \Theta} \min \{ BetP(x), BetP(x') \}$$

- mass assignment:

$$m_{iso}(A_i) = i \cdot (BetP(x_i) - BetP(x_{i+1}))$$

where $\{x_i\} = A_i \setminus A_{i-1}$

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A simple Monte-Carlo approach

to Dempster's combination - Wilson, 1989

- we seek $Bel = Bel_1 \oplus \dots \oplus Bel_m$ on Ω , where the evidence is induced by probability distributions P_i on \mathcal{C}_i via $\Gamma_i : \mathcal{C}_i \rightarrow 2^\Omega$
- Monte-Carlo algorithm **simulates the random set interpretation** of belief functions: $Bel(A) = P(\Gamma(c) \subseteq A | \Gamma(c) \neq \emptyset)$

```

for a large number of trials  $n = 1 : N$  do
  randomly pick  $c \in \mathcal{C}$  such that  $\Gamma(c) \neq \emptyset$ 
  for  $i = 1 : m$  do
    randomly pick an element  $c_i$  of  $\mathcal{C}_i$  with probability  $P_i(c_i)$ 
  end for
  let  $c = (c_1, \dots, c_m)$ 
  if  $\Gamma(c) = \emptyset$  then
    restart trial
  end if
  if  $\Gamma(c) \subseteq A$  then
    trial succeeds,  $T = 1$ 
  end if
end for

```

A Monte-Carlo approach

Wilson, 1989

- the proportion of trials which succeed converges to $Bel(A)$:
 $E[\bar{T}] = Bel(A)$, $Var[\bar{T}] \leq \frac{1}{4N}$
- we say algorithm has accuracy k if $3\sigma[\bar{T}] \leq k$
- picking $c \in \mathcal{C}$ involves m random numbers so it takes $A \cdot m$, A constant
- testing if $x_j \in \Gamma(c)$ takes less than Bm , constant B
- expected time of the algorithm is

$$\frac{N}{1 - \kappa} m \cdot (A + B|\Omega|)$$

where κ is Shafer's conflict measure

- expected time to achieve accuracy k is then $\frac{9}{4(1-\kappa)\kappa^2} m \cdot (A + C|\Omega|)$ for constant C , better for simple support functions
- conclusion: **unless κ is close to 1 (highly conflicting evidence) Dempster's combination is feasible for large values of m (number of BFs to combine) and large Ω (hypothesis space)**

Markov-Chain Monte-Carlo

Wilson and Moral, 1996

- trials are not independent but form a Markov chain
- non-deterministic $OPERATION_i$: changes at most the i -th coordinate $c'(i)$ of c' to y , with chance $P_i(y)$

$$Pr(OPERATION_i(c') = c) \propto P_i(c(i)) \text{ if } c(i) = c'(i), 0 \text{ otherwise}$$

- MCMC algorithm which returns a value $BEL^N(c_0)$ which is the proportion of time in which $\Gamma(c_c) \subseteq X$

```

 $c_c = c_0$ 
 $S = 0$ 
for  $n = 1 : N$  do
  for  $i = 1 : m$  do
     $c_c = OPERATION_i(c_c)$ 
    if  $\Gamma(c_c) \subseteq X$  then
       $S = S + 1$ 
    end if
  end for
end for
return  $\frac{S}{Nm}$ 

```

Importance sampling

Wilson and Moral, 1996

Theorem

If \mathcal{C} is connected (i.e., any c, c' are linked by a chain of $OPERATION_i$) then given ϵ, δ there exist K', N' s.t. for all $K \geq K'$ and $N \geq N'$ and c_0 :

$$Pr(|BEL_K^N(c_0)| < \epsilon) \geq 1 - \delta$$

- further step: **importance sampling** -> pick samples c^1, \dots, c^N according to an “easy to handle” probability distribution P^*
- assign to each sample a weight $w_i = \frac{P(c)}{P^*(c)}$
- if $P(c) > 0$ implies $P^*(c) > 0$ then the average $\frac{\sum_{r(c^j) \subseteq X} w_j}{N}$ is an unbiased estimator of $Bel(X)$
- try to use P^* as close as possible to the real one
- strategies are proposed to compute $P(\mathcal{C}) = \sum_c P(c)$

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Graphical models for belief functions

- to tackle complexity, a number of local computation schemes have been proposed
 - **Barnett's computational scheme**
 - Gordon and Shortliffe's **diagnostic trees**
 - Shafer and Logan's hierarchical evidence
 - **Shafer-Shenoy architecture**
- later on, these developed into graphical models for reasoning with conditional belief functions:
 - Cano et al - propagating uncertainty in directed acyclic networks
 - Xu and Smets - Evidential networks with conditional belief functions
 - Shenoy - graphical representation of valuation-based systems (VBS), called valuation networks
 - Ben Yaghlane and Mellouli - **Directed Evidential Networks**

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Barnett's scheme, 1981

- computations **linear in the size of Ω** if all BFs to combine are simple support focused on singletons or their complements
- simple support function \rightarrow as focal elements only A or Ω
- assume we have a Bel_ω with as FEs only $\{\omega, \bar{\omega}, \Omega\}$ for all ω , and we want to combine them
- uses the fact that the plausibility of the combined BF is a function of their input BFs' **commonalities** $Q(A) = \sum_{B \supseteq A} m(B)$:

$$Pl(A) = \sum_{B \subseteq A, B \neq \emptyset} (-1)^{|B|+1} \prod_{\omega \in \Omega} Q_\omega(B)$$

- we get that $Pl(A) = K \left(1 + \sum_{\omega \in A} \frac{Bel_\omega(\omega)}{1 - Bel_\omega(\omega)} - \prod_{\omega \in A} \frac{Bel_\omega(\bar{\omega})}{1 - Bel_\omega(\omega)} \right)$
- the computation of a specific plausibility value $Pl(A)$ is linear in the size of Ω (only elements of A and not subsets are involved)
- however, the number of events A themselves is still exponential

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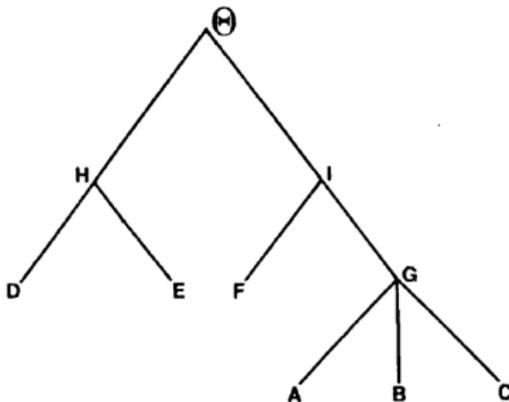
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Gordon and Shortliffe's scheme

based on diagnostic trees

- they are interested in computing degrees of belief **only for events forming a hierarchy (diagnostic tree)**
- (in some applications certain events are not relevant, e.g. classes of diseases)



- combine simple support functions focused on or against the nodes
- produces good approximations, unless evidence is highly conflicting

Gordon and Shortliffe's scheme

based on diagnostic trees

- however, intersection of complements produces FEs not in the tree
- approximated algorithm:
 - 1 first we combine all simple functions focussing on the node events (by Dempster's rule)
 - 2 then, we successively (working down the tree) combine those focused on the complements of the nodes
 - 3 tricky bit: when we do that, we replace each intersection of FEs with the smallest node in the tree that contains it
- results depends on the order of the combination in phase 2
- again approximation can be poor, also no degrees of belief are assigned to complements of nodes
- therefore, we cannot compute their plausibilities!

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Shafer-Shenoy architecture

Qualitative Conditional Independence

- uses **qualitative Markov trees**, which generalise both diagnostic trees and causal trees (Pearl). Extend Pearl's idea to BFs
- partitions Ψ_1, \dots, Ψ_n of a frame are **qualitatively conditionally independent** (QCI) given the partition Ψ if

$$P \cap P_1 \cap \dots \cap P_n \neq \emptyset$$

whenever $P \in \Psi$, $P_i \in \Psi_i$ and $P \cap P_i \neq \emptyset$ for all i

- example: $\{\theta_1\} \times \{\theta_2\} \times \Theta_3$ and $\Theta_1 \times \{\theta_2\} \times \{\theta_3\}$ are QCI on $\Theta_1 \times \Theta_2 \times \Theta_3$ given $\Theta_1 \times \{\theta_2\} \times \Theta_3$ for all $\theta_i \in \Theta_i$
- does not involve probability, but only logical independence
- stochastic conditional independence does imply the above
- if two BFs Bel_1 and Bel_2 are carried by partitions Ψ_1, Ψ_2 which are QCI given Ψ then

$$(Bel_1 \oplus Bel_2)_\Psi = (Bel_1)_\Psi \oplus (Bel_2)_\Psi$$

Qualitative Markov trees

Shafer-Shenoy architecture

- given a tree, deleting a node and all incident edges yields a forest - denote the collection of nodes of the j -th subtree by $\alpha_m(j)$

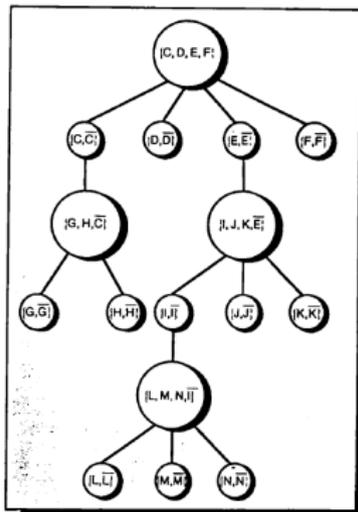


Figure 8. The enlarged qualitative Markov tree for the car that won't start.

- a **qualitative Markov tree** is a tree of partitions s.t. for every node i the minimal refinements of partitions in $\alpha_m(j)$ for $j = 1, \dots, k$ are QCI given Ψ_i
- a Bayesian causal tree becomes a qualitative Markov tree whenever we associate each node B with the partition Ψ_B associated with the random variable v_B
- a QMT remains such if we insert between parent and child their common refinement
- can also be constructed from a diagnostic tree (left), same interpolation property holds

Propagating belief functions

on qualitative Markov trees

- each BF to combine has to be carried by a partition in the tree
- idea: **replace Dempster's combination over the whole frame with multiple implementations over partitions**
- a **processor located at each node** Ψ_i combines BFs using Ψ_i as a frame and projects BFs to its neighbours
 - 1 send Bel_i to its neighbours
 - 2 whenever it gets a new input, computes $(Bel^T)_{\Psi_i} \leftarrow (\oplus \{(Bel_x)_{\Psi_i} : x \in N(i)\} \oplus Bel_i)_{\Psi_i}$
 - 3 computes $Bel_{i,y} \leftarrow (\oplus \{(Bel_x)_{\Psi_i} : x \in N(i) \setminus \{y\}\} \oplus Bel_i)_{\Psi_y}$ and sends it to its neighbour y , for each neighbour
- inputting new BFs in the tree can take place asynchronously
- final result of each processor: **coarsening to that partition of the combination of all inputted BFs:** $(\oplus_{j \in J} Bel_j)_{\Psi_i}$

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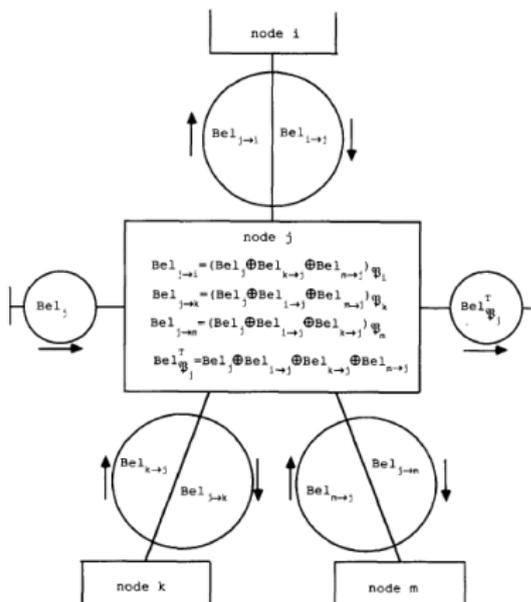


Figure 9. A Typical Processor (with Three Neighbors)

- total time to reach equilibrium is proportional to the tree's diameter

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Directed evidential networks

Ben Yaghlane and Mellouli, 2008

- Evidential networks with conditional belief functions (ENC) were originally proposed by Xu and Smets for the propagation of beliefs
 - (Dempster's conditioning is adopted)
- ENCs contain a directed acyclic graph with conditional beliefs defined in a different manner from conditional probabilities in Bayesian networks (BNs)
 - edges represent the existence of a conditional BF (no form of independence assumed)
 - initially defined only for binary (conditional) relationships
- Ben Yaghlane and Mellouli generalised ENCs to any number of nodes - directed evidential network (DEVN)
 - a directed acyclic graph (DAG) in which directed arcs describe the conditional dependence relations expressed by conditional BFs for each node given its parents
 - new observations introduced in the network are represented by belief functions allocated to some nodes

Directed evidential networks

Ben Yaghlane and Mellouli, 2008

- problem: given n BF's Bel_1, \dots, Bel_n over X_1, \dots, X_n we seek the marginal on X_i of their joint belief function
- uses the generalised Bayesian theorem (GBT) to compute the posterior $Bel(x|y)$ given the conditional $Bel(y|x)$
- the marginal is computed for each node by combining all the messages received from its neighbors and its own prior belief:

$$Bel^X = Bel_0^X \oplus Bel_{Y \rightarrow X}, \quad Bel_{Y \rightarrow X}(x) = \sum_{y \subseteq \Theta_Y} m_0(y) Bel(x|y)$$

where $Bel(x|y)$ is given by GBT

- another application of the message-passing idea to belief functions
- propose a simplified scheme for simply directed networks
- extension to DEVNs by first transforming them to binary join trees

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Decision making with belief functions

An overview

- natural application of belief function representation of uncertainty
- problem: **selecting an act f from an available list \mathcal{F}** (making a “decision”), which optimises a certain objective function
- various approaches to decision making
 - decision making in the TBM is based on expected utility via pignistic transform
 - Strat has proposed something similar in his “cloaked carnival wheel” scenario
 - **generalised expected utility** [Gilboa] based on classical expected utility theory [Savage, von Neumann]
- a lot of interest in **multicriteria decision making** (based on a number of attributes)

Expected utility approach

Decision making under uncertainty

- a decision problem can be formalized by defining:
 - a set Ω of **states of the world**;
 - a set \mathcal{X} of **consequences**;
 - a set \mathcal{F} of **acts**, where an act is a function $f : \Omega \rightarrow \mathcal{X}$
- let \succsim be a **preference relation** on \mathcal{F} , such that $f \succsim g$ means that f is at least as desirable as g
- Savage (1954) has showed that \succsim verifies some rationality requirements iff there exists a **probability measure** P on Ω and a **utility function** $u : \mathcal{X} \rightarrow \mathbb{R}$ s.t.

$$\forall f, g \in \mathcal{F}, \quad f \succsim g \Leftrightarrow \mathbb{E}_P(u \circ f) \geq \mathbb{E}_P(u \circ g)$$

where \mathbb{E}_P denotes the expectation w.r.t. P

- P and u are unique up to a positive affine transformation
- does that mean that basing decisions on belief functions is irrational?

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Decision making in the TBM

Expected utility using the pignistic probability

- in the TBM, decision making is done by **maximising the expected utility of actions based on the pignistic transform**
- (as opposed to computing upper and lower expected utilities directly from (Bel, Pl) via Choquet integral, as we will see later)
- the set of possible actions \mathcal{F} and the set Ω of possible outcomes are distinct, and the utility function is defined on $\mathcal{F} \times \Omega$
- Smets proves the necessity of the pignistic transform by maximizing

$$E[u] = \sum_{\omega \in \Omega} u(f, \omega) Pign(\omega)$$

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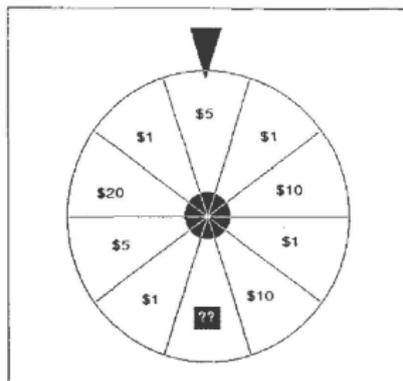
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Strat's decision apparatus

[UAI 1990]



- Strat's decision apparatus is based on computing intervals of expected values
- assumes that the decision frame Ω is itself a set of scalar values (e.g. dollar values, see left) - does not distinguish between utilities and elements of Ω (returns)

- .. so that an **expected value interval** can be computed:
 $E(\Omega) = [E_*(\Omega), E^*(\Omega)]$, where

$$E_*(\Omega) \doteq \sum_{A \subseteq \Omega} \inf(A)m(A), \quad E^*(\Omega) \doteq \sum_{A \subseteq \Omega} \sup(A)m(A)$$

- not good enough to make a decision, e.g.: should we pay a 6\$ ticket when the expected interval is [5\$, 8\$]?

Strat's decision apparatus

A probability of favourable outcome

- Strat identifies ρ as the probability that the value assigned to the hidden sector is the one the player would choose
- $1 - \rho$ is the probability that the sector is chosen by the carnival hawker

Theorem

The expected value of the mass function of the wheel is

$$E(\Omega) = E_*(\Omega) + \rho(E^*(\Omega) - E_*(\Omega))$$

- to decide whether to play the game we only need to assess ρ
- basically, this amounts to a specific probability transform (like the pignistic one)
- Lesh, 1986 had also proposed a similar approach

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Savage's axioms

- Savage has proposed 7 axioms, 4 of which are considered as meaningful (the others are rather technical)
- let us examine the first two axioms:
- Axiom 1: \succsim is a total preorder (complete, reflexive and transitive)
- Axiom 2 [**Sure Thing Principle**]. Given $f, h \in \mathcal{F}$ and $E \subseteq \Omega$, let fEh denote the act defined by

$$(fEh)(\omega) = \begin{cases} f(\omega) & \text{if } \omega \in E \\ h(\omega) & \text{if } \omega \notin E \end{cases}$$

- then the Sure Thing Principle states that $\forall E, \forall f, g, h, h'$,

$$fEh \succsim gEh \Rightarrow fEh' \succsim gEh'$$

- this axiom seems reasonable, but it is **not verified empirically!**

Ellsberg's paradox

- suppose you have an urn containing 30 red balls and 60 balls, either black or yellow. Consider the following gambles:
 - f_1 : you receive 100 euros if you draw a **red ball**
 - f_2 : you receive 100 euros if you draw a **black ball**
 - f_3 : you receive 100 euros if you draw a **red or yellow ball**
 - f_4 : you receive 100 euros if you draw a **black or yellow ball**
- in this example $\Omega = \{R, B, Y\}$, $f_i : \Omega \rightarrow \mathbb{R}$ and $\mathcal{X} = \mathbb{R}$
- the four acts are the mappings in the left table
- empirically it is observed that most people strictly prefer f_1 to f_2 , but they strictly prefer f_4 to f_3

	R	B	Y
f_1	100	0	0
f_2	0	100	0
f_3	100	0	100
f_4	0	100	100

Now, pick $E = \{R, B\}$: by definition

$$f_1\{R, B\}0 = f_1, \quad f_2\{R, B\}0 = f_2$$

$$f_1\{R, B\}100 = f_3, \quad f_2\{R, B\}100 = f_4$$

- since $f_1 \succ f_2$, i.e. $f_1\{R, B\}0 \succ f_2\{R, B\}0$ the Sure Thing principle would imply $f_1\{R, B\}100 \succ f_2\{R, B\}100$, i.e., $f_3 \succ f_4$
- empirically **the Sure Thing Principle is violated!**

Gilboa's theorem

- Gilboa (1987) proposed a modification of Savage's axioms with, in particular, a **weaker form of Axiom 2**
- a preference relation \succsim meets these weaker requirements iff there exists a **(non necessarily additive) measure** μ and a **utility function** $u : \mathcal{X} \rightarrow \mathbb{R}$ such that

$$\forall f, g \in \mathcal{F}, \quad f \succsim g \Leftrightarrow C_\mu(u \circ f) \geq C_\mu(u \circ g),$$

where C_μ is the **Choquet integral**, defined for $X : \Omega \rightarrow \mathbb{R}$ as

$$C_\mu(X) = \int_0^{+\infty} \mu(X(\omega) \geq t) dt + \int_{-\infty}^0 [\mu(X(\omega) \geq t) - 1] dt.$$

- given a belief function Bel on Ω and a utility function u , this theorem supports **making decisions based on the Choquet integral** of u with respect to Bel or Pl

Lower and upper expected utilities

- for finite Ω , it can be shown that

$$C_{Bel}(u \circ f) = \sum_{B \subseteq \Omega} m(B) \min_{\omega \in B} u(f(\omega))$$

$$C_{Pl}(u \circ f) = \sum_{B \subseteq \Omega} m(B) \max_{\omega \in B} u(f(\omega))$$

- let $\mathcal{P}(Bel)$ as usual be the set of probability measures P compatible with Bel , i.e., such that $Bel \leq P$. Then, it can be shown that

$$C_{Bel}(u \circ f) = \min_{P \in \mathcal{P}(Bel)} \mathbb{E}_P(u \circ f) = \underline{\mathbb{E}}(u \circ f)$$

$$C_{Pl}(u \circ f) = \max_{P \in \mathcal{P}(Bel)} \mathbb{E}_P(u \circ f) = \overline{\mathbb{E}}(u \circ f)$$

Decision making

Strategies

- for each act f we have two expected utilities $\underline{\mathbb{E}}(f)$ and $\overline{\mathbb{E}}(f)$. How do we make a decision?
- possible decision criteria based on interval dominance:

- 1 $f \succcurlyeq g$ iff $\underline{\mathbb{E}}(u \circ f) \geq \overline{\mathbb{E}}(u \circ g)$ (**conservative** strategy)

- 2 $f \succcurlyeq g$ iff $\underline{\underline{\mathbb{E}}}(u \circ f) \geq \underline{\underline{\mathbb{E}}}(u \circ g)$ (**pessimistic** strategy)

- 3 $f \succcurlyeq g$ iff $\overline{\overline{\mathbb{E}}}(u \circ f) \geq \overline{\overline{\mathbb{E}}}(u \circ g)$ (**optimistic** strategy)

- 4 $f \succcurlyeq g$ iff

$$\alpha \underline{\mathbb{E}}(u \circ f) + (1 - \alpha) \overline{\mathbb{E}}(u \circ f) \geq \alpha \underline{\mathbb{E}}(u \circ g) + (1 - \alpha) \overline{\mathbb{E}}(u \circ g)$$

for some $\alpha \in [0, 1]$ called a pessimism index (**Hurwicz criterion**)

- the conservative strategy yields only a partial preorder: f and g are not comparable if $\underline{\mathbb{E}}(u \circ f) < \overline{\mathbb{E}}(u \circ g)$ and $\underline{\mathbb{E}}(u \circ g) < \overline{\mathbb{E}}(u \circ f)$

Ellsberg's paradox revisited

- going back to the example, the evidence naturally translates into a belief function
- we have $m(\{R\}) = 1/3$, $m(\{B, Y\}) = 2/3$
- we can then compute lower and upper expected utilities for each action:

	R	B	Y	$\underline{\mathbb{E}}(u \circ f)$	$\overline{\mathbb{E}}(u \circ f)$
f_1	100	0	0	$u(100)/3$	$u(100)/3$
f_2	0	100	0	$u(0)$	$u(200)/3$
f_3	100	0	100	$u(100)/3$	$u(100)$
f_4	0	100	100	$u(200)/3$	$u(200)/3$

- the observed behavior ($f_1 \succcurlyeq f_2$ and $f_4 \succcurlyeq f_3$) is explained by the pessimistic strategy

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Continuous formulations

of the theory of belief functions

- in the original formulation by Shafer [1976], belief functions are defined on finite sets only
- need for generalising this to arbitrary domains has been recognised at an early stage
- main approaches to continuous formulation presented here:
 - Shafer's **allocations of probability** [1982]
 - belief functions as **random sets** [Nguyen]
 - **belief functions on Borel intervals** of the real line [Strat,Smets]
- other approaches, with limited (so far) impact
 - generalised evidence theory
 - MV algebras
 - several others

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Allocations of probability

Shafer, 1979

- every belief function can be represented as an **allocation of probability**, i.e., \cap -homomorphisms into positive and completely additive probability algebra (deduced from the integral representation due to Choquet)
 - for every belief function Bel defined on a class of events $\mathcal{E} \subseteq 2^\Omega$ there exists a complete Boolean algebra \mathcal{M} , a positive measure μ and an allocation of probability ρ between \mathcal{E} and \mathcal{M} such that $Bel = \mu \circ \rho$
- two regularity conditions for a belief function over an infinite domain are considered: **continuity** and **condensability**
- canonical continuous extensions** of belief functions to arbitrary power sets can be introduced by allocation of probability
- the approach shows significant resemblance with the notions of inner measure and extension of capacities [Honda]

Continuity and condensability

Shafer's allocations of probability

- $\mathcal{E} \subset 2^\Theta$ is a multiplicative subclass of 2^Θ if $A \cap B \in \mathcal{E}$ for all $A, B \in \mathcal{E}$
- a function $Bel : \mathcal{E} \rightarrow [0, 1]$ such that $Bel(\emptyset) = 0$, $Bel(\Theta) = 1$ and Bel is monotone of order ∞ is a belief function
 - equally, an upper probability (plausibility) function is alternating of order ∞ (\geq is exchanged with \leq)
- a BF on 2^Θ is **continuous** if $Bel(\bigcap_i A_i) = \lim_{i \rightarrow \infty} Bel(A_i)$ for every decreasing sequence of A_i s. A BF on a multiplicative subclass \mathcal{E} is continuous if it can be extended to a continuous one on 2^Θ
 - continuity arises from partial beliefs on 'objective' probabilities
- a BF on 2^Θ is **condensable** if $Bel(\bigcap \mathcal{A}) = \inf_{A \in \mathcal{A}} Bel(A)$ for every downward net \mathcal{A} in 2^Θ . A BF on a multiplicative subclass \mathcal{E} is condensable if it can be extended to a condensable one on 2^Θ
 - a downward net is such that given two elements there is always an element subset of their intersection
- condensability is restrictive, but related to Dempster's rule

Choquet's representation

Shafer's allocations of probability

- Choquet's integral representation says that every belief function can be represented by allocation of probability
- $r : \mathcal{E} \rightarrow \mathcal{F}$ is a \cap -homomorphism if it preserves \cap

Choquet's theorem

For every BF Bel on a multiplicative subclass \mathcal{E} of 2^Θ , \exists a set \mathcal{X} and an algebra \mathcal{F} of its subsets, a finitely additive probability measure μ on \mathcal{F} , and a \cap -homomorphism $r : \mathcal{E} \rightarrow \mathcal{F}$ such that $Bel = \mu \circ r$.

- if we replace the measure space $(\mathcal{X}, \mathcal{F}, \mu)$ with a probability algebra (a complete Boolean algebra \mathcal{M} with a completely additive prob measure μ) we get

Allocation of probability

For every BF Bel on a multiplicative subclass \mathcal{E} of 2^Θ , \exists an allocation of probability $\rho : \mathcal{E} \rightarrow \mathcal{M}$ such that $Bel = \mu \circ \rho$.

- non-zero elements of \mathcal{M} can be thought of as focal elements

Canonical extension

Shafer's allocations of probability

Theorem

a BF on a multiplicative subclass \mathcal{E} can always be extended to a belief function on 2^Ω by **canonical extension**

$$\overline{Bel}(A) \doteq \sup_{n \geq 1, A_1, \dots, A_n \in \mathcal{E}} \sum \left\{ (-1)^{|I|+1} Bel(\cap_{i \in I} A_i) \mid \emptyset \neq I \subset \{1, \dots, n\} \right\}$$

- proof is based on the existence of an allocation for the extension
- note the similarity with the superadditivity axiom
- also related to **inner measures**, which provide approximate belief values for subsets not in a sigma-algebra
- \overline{Bel} is the minimal such extension
- what about evidence combination? **condensability** ensures that the Boolean algebra \mathcal{M} represents intersection properly for arbitrary (not just finite) collections \mathcal{B} of subsets:

$$\rho(\cap \mathcal{B}) = \bigwedge_{B \in \mathcal{B}} \rho(B) \quad \forall \mathcal{B} \subset 2^\Omega$$

- allows us to imagine Dempster's combinations of infinitely many belief functions

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Random sets

to extend belief functions to arbitrary domains

- the notion of condensability has been studied by Nguyen for upper probabilities generated by random sets too [Nguyen 1978]
- efforts directed at a general theory on arbitrary domains
- for finite random sets (i.e. with a finite number of focal elements), under independence of variables **Dempster's rule can be applied:**

$$(\mathcal{F}, m) = \left\{ A_{i_1, \dots, i_d} = \times_{j=1}^d A_{i_j}, m_{i_1, \dots, i_d} = m_{i_1} \cdot \dots \cdot m_{i_d} \right\}$$

- for dependent sources Fetz and Oberguggenberger have proposed an “unknown interaction” model
- for infinite random sets Alvarez (see p-boxes later) a **Monte-Carlo sampling** method

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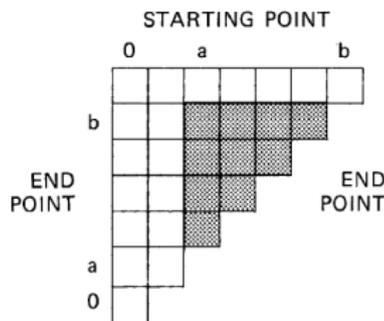
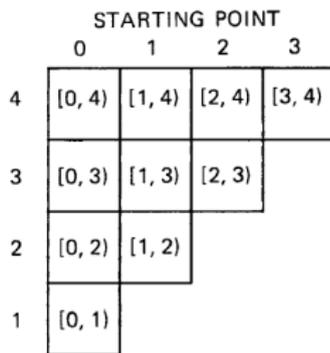
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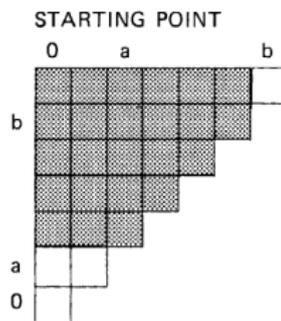
Continuous belief functions

Strat's approach

- idea: take a real interval I and split it into N bits
- take as frame of discernment the **set of possible intervals** with these extreme: $[0, 1)$, $[0, 2)$, $[1, 4)$ etc
- a belief function there has $\sim N^2/2$ possible focal elements, so that its mass lives on a triangle (left), and one can compute **belief and plausibility by integration** (right)



(a) $Spt([a, b])$

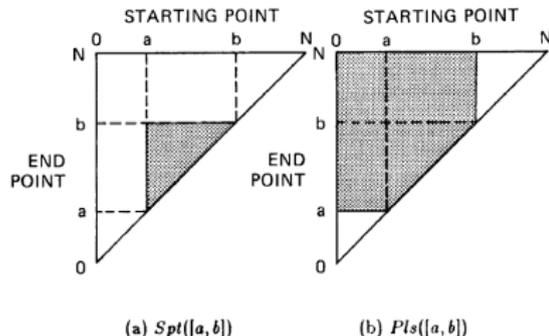
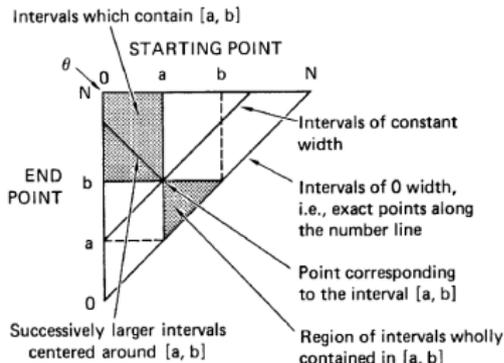


(b) $Pls([a, b])$

Continuous belief functions

Strat's approach

- this trivially generalises to **all arbitrary intervals of I** (below)



$$Bel([a, b]) = \int_a^b \int_x^b m(x, y) dy dx, \quad Pl([a, b]) = \int_0^b \int_{\max(a, x)}^N m(x, y) dy dx$$

- Dempster's rule generalises as $Bel_1 \oplus Bel_2([a, b]) =$**

$$\frac{1}{K} \int_0^a \int_b^N [m_1(x, b)m_2(a, y) + m_2(x, b)m_1(a, y) + m_1(a, b)m_2(x, y) + m_2(a, b)m_1(x, y)] dy dx$$

Continuous belief functions

on the Borel algebra of intervals

- a pretty much identical approach is followed by Smets
- allows us to define a **continuous pignistic PDF** as

$$Bet(a) \doteq \lim_{\epsilon \rightarrow 0} \int_0^a dx \int_{a+\epsilon}^1 \frac{m(x, y)}{y-x} dy$$

- can be easily extended to the real line, by considering **belief functions defined on the Borel σ -algebra** of subsets of \mathbb{R} generated by the collection \mathcal{I} of closed intervals
- the theory provides a way of **building a continuous belief function from a pignistic density**, by applying the least commitment principle and assuming unimodal pignistic PDFs

$$Bel(s) = -(s - \bar{s}) \frac{dBet(s)}{ds}$$

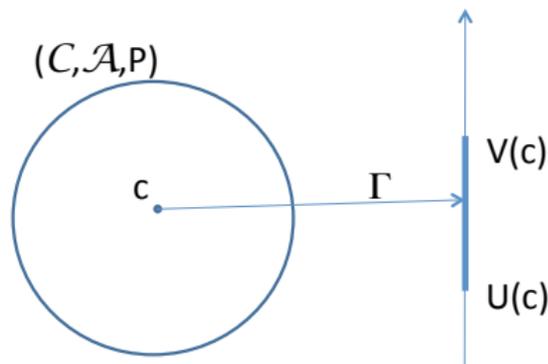
where \bar{s} is such that $Bet(s) = Bet(\bar{s})$

- example: $Bet(x) = \mathcal{N}(x, \mu, \sigma)$ is normal $\rightarrow Bel(y) = \frac{2y}{\sqrt{2\pi}} e^{-y^2}$,
where $y = (x - \mu)/\sigma$

Continuous belief functions

induced by random closed intervals

- formal setting:
- let (U, V) be a two-dimensional random variable from $(\mathcal{C}, \mathcal{A}, P)$ to $(\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$ such that $P(U \leq V) = 1$ and $\Gamma(c) = [U(c), V(c)] \subseteq \mathbb{R}$

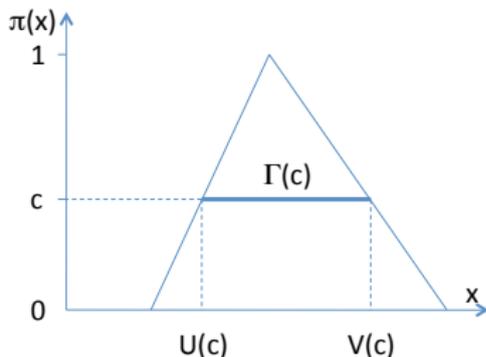


- this setting defines a **random closed interval**, which induces a belief function on $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$ defined by

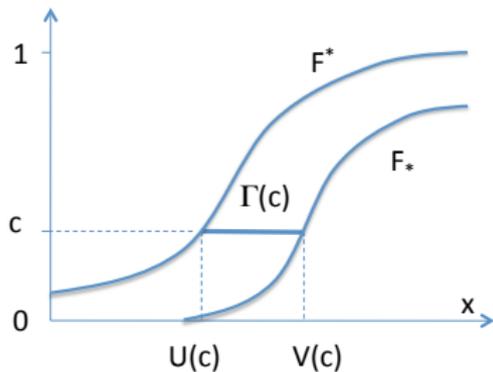
$$Bel(A) = P([U, V] \subseteq A), \quad \forall A \in \mathcal{B}(\mathbb{R})$$

Special cases of random closed intervals

Consonant random interval



p-box



- special cases
- a fuzzy set on the real line induces a mapping to a collection of nested intervals, parameterised by the level c
- a p-box, i.e, upper and lower bounds to a cumulative distribution function (see later) also induces a family of intervals

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T. Denoeux

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Relationships with other theories of uncertainty

- belief functions have meaningful relationships with a number of other theories of uncertainty
- here we briefly recall the most significant ones:
 - **imprecise probabilities** [Walley]
 - **credal sets** [Levi]
 - **possibility theory** [Zadeh, Dubois & Prade]
 - belief functions **on fuzzy sets** [Zadeh & others]
 - **p-boxes** [Ferson]
- others we will not touch here for lack of time:
 - probability intervals [Moral]
 - monotone capacities
 - fuzzy measures
 - rough sets [Pawlak]
 - probabilistic and modal logic

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Lower probabilities and credal sets

in Imprecise Probability

- a **lower probability** \underline{P} is a function from 2^Ω , the power set of Ω , to the unit interval $[0, 1]$
- with any lower probability \underline{P} is associated a dual upper probability function \overline{P} , defined for any $A \subseteq \Omega$ as $\overline{P}(A) = 1 - \underline{P}(A^c)$
- with any lower probability \underline{P} we can associate a closed convex set (**credal set** [Levi])

$$\mathcal{P}(\underline{P}) = \left\{ P : P(A) \geq \underline{P}(A), \forall A \subseteq \Omega \right\}$$

of probability measures P which dominate \underline{P}

- note that not all convex sets of probabilities can be described by merely focusing on events [Walley]

Coherent lower probabilities

in Imprecise Probability

- a lower probability \underline{P} is called 'consistent' if $\mathcal{P}(\underline{P}) \neq \emptyset$ and 'tight' if

$$\inf_{p \in \mathcal{P}(\underline{P})} P(A) = \underline{P}(A)$$

- (respectively \underline{P} 'avoids sure loss' and \underline{P} is 'coherent' in Walley's terminology)
- consistency means that the lower bound constraints $\underline{P}(A)$ can indeed be satisfied by some probability measure
- tightness indicates that \underline{P} is the lower envelope on subsets of $\mathcal{P}(\underline{P})$
- belief functions are indeed a **special type of coherent lower probabilities**, which in turn can be seen as a special class of *lower previsions*
- having said that, the two approaches depart on the fundamental epistemic representation of evidence

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Possibility theory

and consonant belief functions

- a **possibility measure** on a domain Ω is a function $Pos : 2^\Omega \rightarrow [0, 1]$ such that $Pos(\emptyset) = 0$, $Pos(\Omega) = 1$ and

$$Pos \left(\bigcup_i A_i \right) = \sup_i Pos(A_i)$$

for any family $\{A_i | A_i \in 2^\Omega, i \in I\}$ where I is an arbitrary set index

- it is uniquely characterized by a **membership function** or “possibility distribution” $\pi(x) \doteq Pos(\{x\})$, as $Pos(A) = \sup_{x \in A} \pi(x)$
 - $Nec(A) = 1 - Pos(A^c)$ is called **necessity measure**
- call “plausibility assignment” pl the restriction of the plausibility function to singletons $pl(x) = Pl(\{x\})$ - then [Shafer]:
- **Bel is a necessity measure iff Bel is consonant**
- the membership function coincides with the plausibility assignment
 - according to Shafer, the difference between possibilities and consonant BFs is just in the language used

Belief functions on fuzzy sets

- 1 a finite fuzzy set is equivalent to a consonant belief function
- 2 **generalisations of belief functions defined on fuzzy sets** have been proposed [Zadeh]

- basic idea: belief measures generalised on fuzzy sets as follows:

$$Bel(X) = \sum_{A \in \mathcal{M}} I(A \subseteq X) m(A)$$

where X is a fuzzy set defined on Ω , m is a mass function defined on the collection of fuzzy sets on Ω

- $I(A \subseteq X)$ is a measure of how much the fuzzy set A is included in the fuzzy set X
- various measures of inclusion in $[0, 1]$ can be proposed:
 - Lukasiewicz: $I(x, y) = \min\{1, 1 - x - y\}$ [Ishizuka]
 - Kleene-Dienes: $I(x, y) = \max\{1 - x, y\}$ [Yager]
- from which one can get: $\mathcal{I}(A, B) = \bigwedge_{x \in \Theta} I(A(x), B(y))$ [Wu 2009]

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Probability-boxes (p-Boxes)

Classes of cumulative distribution functions

- a **probability box** or **p-box** [Ferson and Hajagos] $\langle \underline{F}, \overline{F} \rangle$ is a class of cumulative distribution functions (CDFs)

$$\langle \underline{F}, \overline{F} \rangle = \{ F \text{ CDF} : \underline{F} \leq F \leq \overline{F} \}$$

delimited by upper and lower CDF bounds \underline{F} and \overline{F}

- represents the epistemic uncertainty about the CDF of a random variable
- every RS generates a unique p-box whose CDFs are all those consistent with the evidence:

$$\underline{F}(x) = Bel((-\infty, x]), \quad \overline{F}(x) = Pl((-\infty, x])$$

- every p-box generates an infinite RS with as focal elements the following infinite collection of intervals of \mathbb{R} :

$$\{ [\overline{F}^{-1}(\alpha), \underline{F}^{-1}(\alpha)] \mid \forall \alpha \in [0, 1] \}$$

where $\overline{F}^{-1}(\alpha) \doteq \inf\{\overline{F}(x) \geq \alpha\}$, $\underline{F}^{-1}(\alpha) \doteq \inf\{\underline{F}(x) \geq \alpha\}$

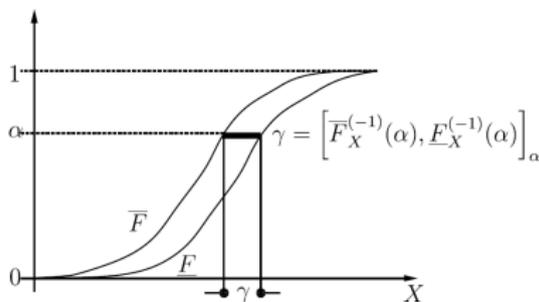
Approximate computations

for random sets

- in an infinite RS the computation of the integral $Bel(A) = \int_{c \in \mathcal{C}} I[\Gamma(c) \subset A] dP(c)$ (or those for $Pl(A)$, etc) is not trivial
- we can use the representation of infinite RSs provided by p-boxes, with set of focal elements

$$\mathcal{F} = \left\{ \gamma = [\bar{F}^{-1}(\alpha), \underline{F}^{-1}(\alpha)] \forall \alpha \in [0, 1] \right\}$$

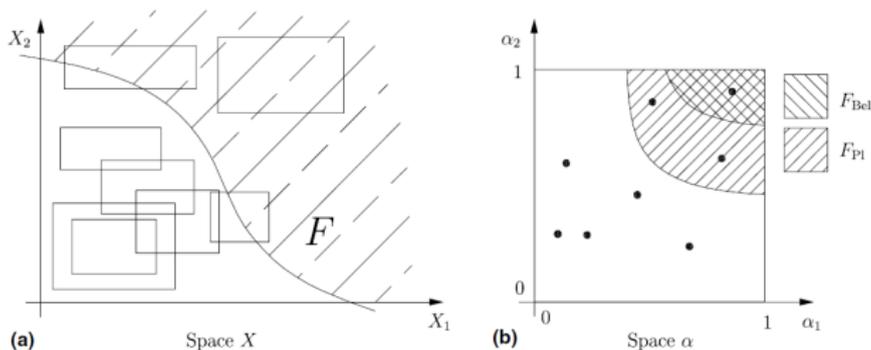
- if α has its own CDF F_α , we can sample from it
- after sampling FEs from the RS, we can compute belief and plausibility integrals



Approximate combination of random sets

α -representation

- we can also calculate the combination of the sampled FEs
- if d random sets to combine, FEs are vectors of indices from all constituting RS: $\alpha = [\alpha_1, \dots, \alpha_d] \in (0, 1]^d$



- suppose a copula C is defined on the unit hypercube (i.e. a prob distribution whose marginals are uniform)..

Approximate combination of random sets

Monte-Carlo approach

- ..we can use it to compute the desired integrals, i.e.

$$P_{\Gamma}(G) = \int_{\alpha \in G} dC(\alpha)$$

- if input RS are independent, these integrals decompose as, e.g.

$$\text{Bel}_{(\mathcal{F}, P_{\Gamma})}(F) = \underbrace{\int_{0^+}^1 \cdots \int_{0^+}^1}_{d\text{-times}} I[[\alpha_1, \dots, \alpha_d] \in F_{\text{Bel}}] dC(\alpha_1, \dots, \alpha_d)$$

- Monte-Carlo approach [Alvarez 2006] - for $j = 1, \dots, n$:
 - randomly extract a sample α_j from the copula C
 - form the corresponding focal element $A_j = \times_{i=1, \dots, d} \gamma_i^d$
 - assign to it mass $m(A_j) = \frac{1}{n}$
- can prove that such approximation converges as $n \rightarrow +\infty$ almost surely to the actual random set

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The mathematics of belief functions

- belief functions are rather complex mathematical objects, therefore:
 - have links with a number of fields of (applied) mathematics
 - lead to interesting generalisations of standard results of classical probability (e.g. Bayes' theorem, total probability)
- **matrix representation**
- **geometric approach** to uncertainty [Cuzzolin]
- measuring **distances** [Jousselme et al]
- algebra of frames [Kohlas]
- abstract independence, Boolean algebras and matroids [Cuzzolin]
- **Moebius transforms**
- entropy and other **measures of uncertainty** [Yager, Klir, Harmanec]

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Matrix representation

Linear algebra

- given an ordering of the subsets of Ω mass, belief, and plausibility functions can be represented as vectors \mathbf{m} , \mathbf{bel} and \mathbf{pl}
- various operations with belief functions can be expressed via vectors and matrices
- negation ($\overline{m}(A) = m(\overline{A})$): $\overline{\mathbf{m}} = \mathbf{Jm}$ where \mathbf{J} is the matrix whose inverse diagonal is made of 1s
- belief value: $\mathbf{b} = \mathbf{BfrMm}$, where $\mathbf{BfrM}(A, B) = 1$ iff $B \subseteq A$ and 0 otherwise
- $\mathbf{BfrM}_{i+1} = \begin{bmatrix} 1 & 0 \\ 1 & 1 \end{bmatrix} \otimes \mathbf{BfrM}_i$ where \otimes is the Kronecker product
- other transformation matrices for Moebius inversion can be defined:

$$\mathbf{MfrB} = \mathbf{BfrM}^{-1}, \quad \mathbf{QfrM} = \mathbf{JBfrMJ}, \quad \mathbf{MfrQ} = \mathbf{JBfrM}^{-1}\mathbf{J}$$

- normalised BFs and plausibilities: $\mathbf{Bel} = \mathbf{b} - b(\emptyset)\mathbf{1}$, $\mathbf{pl} = \mathbf{1} - \mathbf{Jb}$

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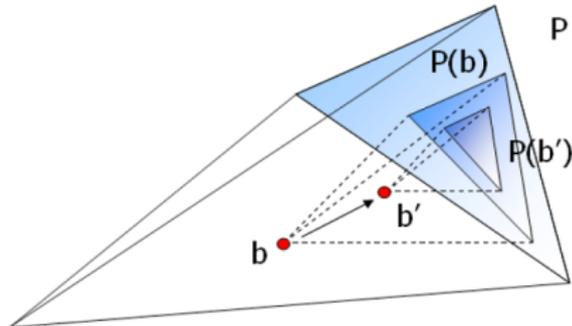
A geometric approach

to the theory of evidence

- the collection \mathcal{B} of all the vectors $\mathbf{b} = [Bel(A), \emptyset \subsetneq A \subsetneq \Omega]'$ representing a belief function on Ω is a “simplex” (in rough words a higher-dimensional triangle), the **belief space**

$$\mathcal{B} = Cl(\mathbf{b}_A, \emptyset \subsetneq A \subseteq \Omega)$$

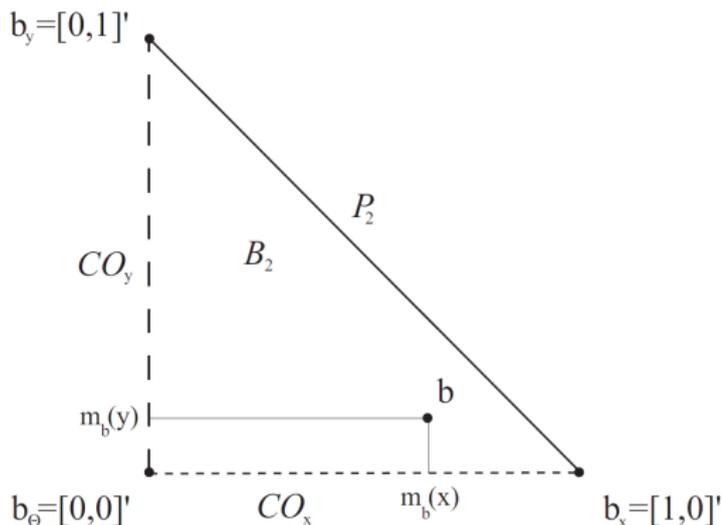
which is the convex closure of (the vectors of) all “logical” BFs \mathbf{b}_A



- alternatively we can adopt mass vectors $\mathbf{m}_b = [m_b(A), \emptyset \subsetneq A \subseteq \Omega]'$, living in a **mass space**: $\mathcal{M} = Cl(\mathbf{m}_A, \emptyset \subsetneq A \subseteq \Omega)$

Binary example

The simplex of BFs on a frame of size 2



- belief/mass space $B_2 = \mathcal{M}_2$ for a binary frame
- set of probabilities is a face of the simplex (triangle)
- region of consonant BFs is a “simplicial complex” $CO = \bigcup_{x \in \Omega} CI(\mathbf{b}_A, A \ni x)$

Geometry of Dempster's rule

Conditional subspaces

- Dempster's rule behavior w.r.t. affine combination

$$\mathbf{b} \oplus \sum_i \alpha_i \mathbf{b}_i = \sum_i \beta_i (\mathbf{b} \oplus \mathbf{b}_i), \quad \beta_i = \frac{\alpha_i \kappa(\mathbf{b}, \mathbf{b}_i)}{\sum_{j=1}^n \alpha_j \kappa(\mathbf{b}, \mathbf{b}_j)}$$

where $\kappa(\mathbf{b}, \mathbf{b}_i)$ is the usual Dempster's conflict

- convex closure (Cl) and \oplus **commute** in the belief space

$$\mathbf{b} \oplus Cl(\mathbf{b}_1, \dots, \mathbf{b}_n) = Cl(\mathbf{b} \oplus \mathbf{b}_1, \dots, \mathbf{b} \oplus \mathbf{b}_n)$$

- the **conditional subspace** $\langle \mathbf{b} \rangle$ - the set of all BFs (Dempster-) conditioned by \mathbf{b} :

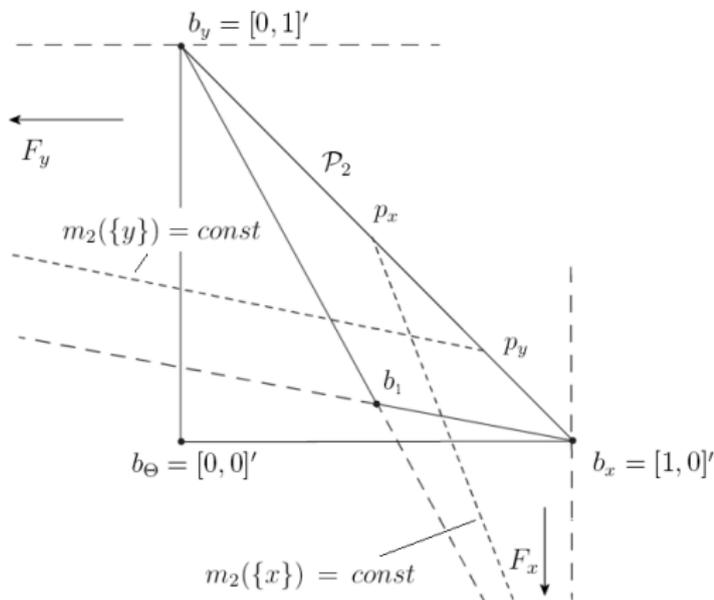
$$\langle \mathbf{b} \rangle \doteq \left\{ \mathbf{b} \oplus \mathbf{b}', \forall \mathbf{b}' \in \mathcal{B} \text{ s.t. } \exists \mathbf{b} \oplus \mathbf{b}' \right\}$$

is the convex closure

$$\langle \mathbf{b} \rangle = Cl(\mathbf{b} \oplus \mathbf{b}_A, \forall A \subseteq \mathcal{C}_{\mathbf{b}})$$

Geometry of Dempster's rule

Geometric construction



- the pointwise behavior of \oplus depends on the notions of **constant mass locus** [Cuzzolin, 2004] and of **foci** $\{F_x, x \in \Omega\}$ of a conditional subspace

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Dissimilarity measures

between belief functions - an overview

- a number of norms can be introduced for belief functions
- e.g., generalizations to belief functions of the classical Kullback-Leibler divergence of two probability distributions P, Q :

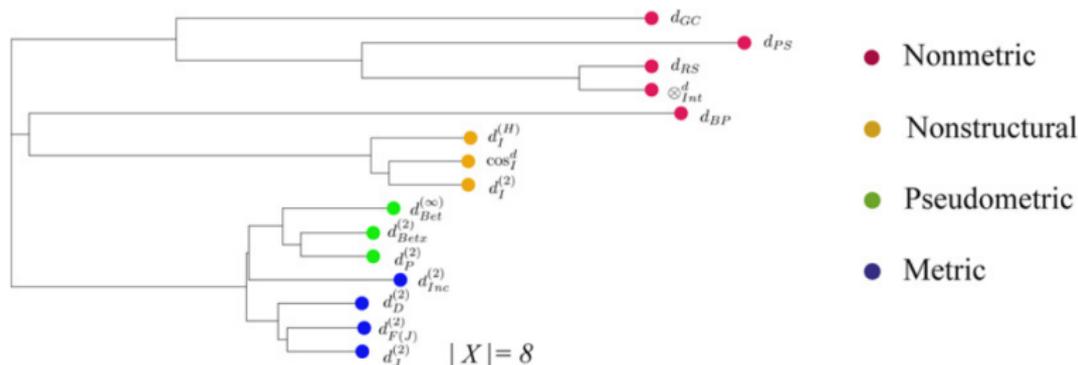
$$D_{KL}(P|Q) = \int_{-\infty}^{\infty} p(x) \log\left(\frac{p(x)}{q(x)}\right) dx$$
- measures based on information theory such as fidelity and entropy-based norms [Jousselme IJAR'11]
- many others have been proposed [diaz,jiang,khatibi,shi], exhaustive analysis huge task!

		Euclidean L_2 (m) Divergence Perry and Stephanou, 1991	Cross-entropy Denoeux, 2000	Information-based Denoeux, 2001 Weighted L_2 (Jaccard) Jousselme <i>et al.</i> , 2001	Cosine similarity Wen <i>et al.</i> , 2008 Euclidean L_2 (Bel) Cuzzolin, 2008		
	1970	...	1990	1995	2000	2005	2010
Conflict Dempster, 1967			Chebyshev L_∞ Tessem, 1993	BPAM Fixsen and Mahler, 1997		2D coefficient Liu, 2006	Euclidean L_p (Bel) Cuzzolin, 2009
				Squared error Zouhal and Denoeux, 1998		Modified weighted L_2 Diaz <i>et al.</i> , 2006	
				Attribute distance Blackman and Popoli, 1999		TBM pairwise dissimilarity Hellinger measure Ristic and Smets, 2006	
				Manhattan L_1 (Bel) Klir, Harmanec, 1999			

Families of distances

between belief functions

- experimental tests on randomly generated BFs lead to the emergence of four families
 - metric (i.e. proper distance functions)
 - pseudo-metric (dissimilarities)
 - non-structural (do not account for structure of focal elements)
 - non-metric



Jousselme's distance

- most popular and cited measure of dissimilarity
- was proposed as a “measure of performance” of algorithms (e.g. object identification) where successive evidence combination leads to convergence to the “true” solution
- based on the geometric representation of mass functions m

$$d_J(m_1, m_2) \doteq \sqrt{\frac{1}{2}(\mathbf{m}_1 - \mathbf{m}_2)^T D(\mathbf{m}_1 - \mathbf{m}_2)}$$

where $D(A, B) = \frac{|A \cap B|}{|A \cup B|}$ for all $A, B \in 2^\Theta$

- D so defined:
 - is definite positive, therefore it defines a metric distance
 - takes into account the similarity among subsets
 - is such that $d(A, B) < d(A, C)$ is C is “closer” to A than B
- this notion remains not well specified though

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Moebius inverses

of plausibilities and commonalities?

- belief function are sum functions: $Bel(A) = \sum_{B \subseteq A} m(B)$
- analogous of integral in calculus, derivative = Moebius inversion
- plausibilities and commonalities **have Moebius inverses**
- only, b.pl.a.s can be negative; b.comm.a.s are not even normalised

belief function

$$m_b(A) = \sum_{B \subseteq A} (-1)^{|A-B|} b(B)$$

b.b.a.

plausibility
function

$$\mu_b(A) \doteq \sum_{B \subseteq A} (-1)^{|A-B|} pl_b(B)$$

b.pl.a.

commonality
function

$$q_b(B) = \sum_{\emptyset \subseteq A \subseteq B} (-1)^{|B \setminus A|} Q_b(A)$$

b.comm.a.

- plausibilities and commonalities **live in simplices congruent with the belief space \mathcal{B}**

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Measures of uncertainty

for belief functions

- various measures have been proposed - see for instance the (rather outdated) survey by Nikhil Pal, 1992
- **Yager's entropy measure** (350+ citations):

$$E(m) = - \sum_{A \in \mathcal{F}} m(A) \log Pl(A)$$

- Yager's entropy is 0 for consonant or **consistent** BFs ($A_i \cap A_j \neq \emptyset$ for all FEs)
 - is maximal for disjoint focal elements with equal mass
- Hohle's measure of confusion: $C(m) = - \sum_{A \in \mathcal{F}} m(A) \log Bel(A)$
- **specificity** of belief measures: $N(m) = \sum_{A \in \mathcal{F}} \frac{m(A)}{|A|}$
 - measure the dispersion of the evidence
 - clearly related to pignistic function
- Klir's non-specificity (extended by Dubois & Prade):

$$I(m) = \sum_{A \in \mathcal{F}} m(A) \log |A|$$

Measures of uncertainty

Global measures

- composite measures: Lamata and Moral: $E(m) + I(m)$
- $E(m)$ was criticised by Klir & Ramer for it expresses conflict as $A \cap B = \emptyset$ rather than $B \not\subseteq A$
- $C(m)$ was criticised for it does not measure to what extent two focal elements disagree (size of $A \cap B$)
- Klir & Ramer's **global uncertainty measure**: $D(m) + I(m)$, where

$$D(m) = - \sum_{A \in \mathcal{F}} m(A) \log \left[\sum_{B \in \mathcal{F}} m(B) \frac{|A \cap B|}{|B|} \right]$$

- Pal argues that none of them is really satisfactory: none of the composite measures have a unique maximum
- there is no sounding rationale for simply adding conflict and non-specificity measures together to get a “total” one
- some are computationally very expensive

Aggregated Uncertainty vs Ambiguity Measure

- Harmanec's **Aggregated Uncertainty** (AU) as the maximal Shannon entropy of all consistent probabilities
 - obviously assumes a credal set interpretation
 - it is the minimal measure meeting eight requirements: symmetry, continuity, expansibility, subadditivity, additivity, monotonicity, normalisation
- criticised by Klir and Smith for being insensitive to changes in evidence
- replaced by a linear combination of AU and nonspecificity $I(m)$
- still high computational complexity
- Josselme et al, 2006: **Ambiguity Measure** (AM), basically classical entropy of pignistic function

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A set of tools for the working scientist

using belief functions

- scientists face on a daily basis problems such as:
 - **making decisions** based on the available data (we already covered this)
 - **estimating** a quantity of interest given the available data (which can be missing, incomplete, conflicting, partially specified)
 - **classifying** data-points into bins
 - extending k-NN classification approaches
 - fusing the results of multiple classifiers
 - **clustering** clouds of data to make sense of them
 - learning a mapping from measurements to a domain of interest (**regression**)
 - ranking objects
- belief functions can provide useful approaches to all these problems when in the presence of (heavy) uncertainty

Belief functions in Machine Learning

- the theory of belief functions has great potential to help solve **complex machine learning (ML) problems**, particularly those involving:
 - **weak information** (partially labeled data, unreliable sensor data, etc.);
 - **multiple sources** of information (classifier or clustering ensembles) [Quost et al., 2007; Masson & Denoeux, 2011]
- other recent ML applications of belief functions:
 - regression [Petit-Renaud & Denoeux, 2004]
 - multi-label classification [Denoeux et al. 2010]
 - clustering [Masson & Denoeux, 2008; Antoine et al., 2012]

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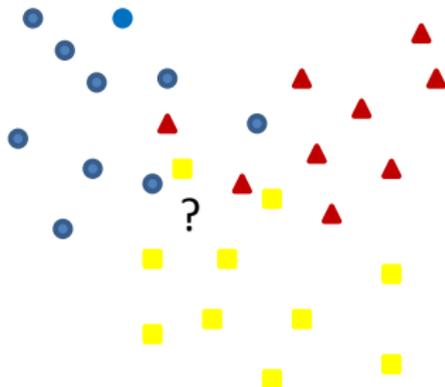
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Classification problems

in machine learning



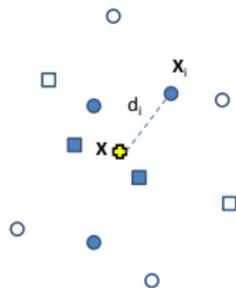
- a population is assumed to be partitioned in c groups or classes
- let $\Omega = \{\omega_1, \dots, \omega_c\}$ denote the set of classes
- each instance is described by
 - a feature vector $\mathbf{x} \in \mathbb{R}^p$
 - a class label $y \in \Omega$
- problem: given a **learning set** $\mathcal{L} = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)\}$, **predict the class** of a new instance described by \mathbf{x}

Main approaches

Classification with belief functions

- 1 Approach 1 (**ensemble classification**): Convert the outputs from standard classifiers into belief functions and combine them using Dempster's rule or any other alternative rule (e.g., Quost al., *IJAR*, 2011)
 - 2 Approach 2: Develop **evidence-theoretic classifiers** directly providing belief functions as outputs:
 - **Generalized Bayes theorem**, extends the Bayesian classifier when class densities and priors are ill-known [Appriou, 1991; Denoëux & Smets, 2008]
 - **Distance-based approach**: evidential k -NN rule [Denoëux, 1995], evidential neural network classifier [Denoëux, 2000]
- today we will just see the evidential k -NN rule (for complete and partially specified data)

Evidential K-NN



- let Ω be the set of classes
- let $\mathcal{N}_k(\mathbf{x}) \subset \mathcal{L}$ denote the set of the k nearest neighbors of \mathbf{x} in \mathcal{L} , based on some distance measure d
- each $\mathbf{x}_i \in \mathcal{N}_k(\mathbf{x})$ can be considered as a **piece of evidence** regarding the class of \mathbf{x} represented by a mass function m_i on Ω :

$$m_i(\{y_i\}) = \varphi(d_i), \quad m_i(\Omega) = 1 - \varphi(d_i)$$

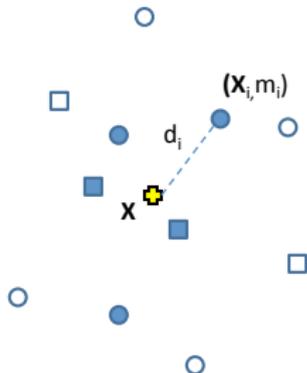
- the **strength of this evidence decreases with the distance** d_i between \mathbf{x} and \mathbf{x}_i - φ is a **decreasing function** such that $\lim_{d \rightarrow +\infty} \varphi(d) = 0$
- pooling of evidence: $m = \bigoplus_{\mathbf{x}_i \in \mathcal{N}_k(\mathbf{x})} m_i$
- the function φ can be fixed heuristically or selected among a family $\{\varphi_\theta | \theta \in \Theta\}$ using, e.g., cross-validation
- decision: select the class with the **highest plausibility**

Classification with partially specified data

- in some applications, learning instances are labeled by experts or indirect methods (no ground truth)
- **class labels of learning data are then uncertain**: partially supervised learning problem
- formalization of the learning set: $\mathcal{L} = \{(\mathbf{x}_i, m_i), i = 1, \dots, n\}$, where
 - \mathbf{x}_i is the attribute vector for instance i , and
 - m_i is a mass function representing **uncertain expert knowledge** about the class y_i of instance i
- special cases:
 - $m_i(\{\omega_k\}) = 1$ for all i : **supervised learning**
 - $m_i(\Omega) = 1$ for all i : **unsupervised learning**
- the evidential k -NN rule can easily be adapted to handle such uncertain learning data

Evidential k -NN rule

for partially supervised data



- Ω is again the collection of classes
- each mass function m_i is **discounted** with a rate depending on the distance d_i :

$$m'_i(A) = \varphi(d_i) m_i(A), \quad \forall A \subset \Omega$$

$$m'_i(\Omega) = 1 - \sum_{A \subset \Omega} m'_i(A)$$

- the k mass functions m'_i are combined using **Dempster's rule**:

$$m = \bigoplus_{\mathbf{x}_i \in \mathcal{N}_k(\mathbf{x})} m'_i$$

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Ranking aggregation

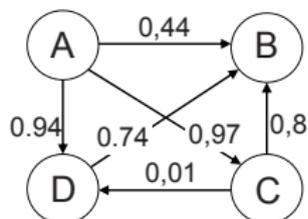
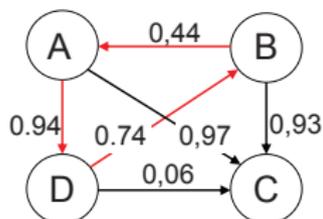
Problem definition

- we consider a **set of alternatives** $O = \{o_1, o_2, \dots, o_n\}$ and an **unknown linear order** (transitive, antisymmetric and complete relation) on O
- typically, this linear order corresponds to **preferences** held by an agent or a group of agents, so that $o_i \succ o_j$ is interpreted as “alternative o_i is preferred to alternative o_j ”
 - (compare inference from qualitative data)
- a source of information (elicitation procedure, classifier) provides us with $n(n-1)/2$ **paired comparisons**, affected by uncertainty
- problem: derive the **most plausible linear order** from this uncertain (and possibly conflicting) information

Example

Tritchler & Lockwood, 1991

- consider four scenarios $O = \{A, B, C, D\}$ describing ethical dilemmas in health care
- suppose two experts gave their preference for all six possible scenario pairs with confidence degrees described below



- assuming the existence of a unique **consensus linear ordering** L^* and seeing the expert assessments as sources of information, what can we say about L^* ?

Combining pairwise masses

on the space of linear orders

- the frame of discernment is the set \mathcal{L} of **linear orders over O**
- comparing each pair of objects (o_i, o_j) yields a **pairwise mass function** $m^{\Theta_{ij}}$ on a coarsening $\Theta_{ij} = \{o_i \succ o_j, o_j \succ o_i\}$ with:

$$m^{\Theta_{ij}}(o_i \succ o_j) = \alpha_{ij}, \quad m^{\Theta_{ij}}(o_j \succ o_i) = \beta_{ij}, \quad m^{\Theta_{ij}}(\Theta_{ij}) = 1 - \alpha_{ij} - \beta_{ij}$$

- $m^{\Theta_{ij}}$ may come from a single expert (e.g., an evidential classifier) or from the combination of the evaluations of several experts
- let $\mathcal{L}_{ij} = \{L \in \mathcal{L} \mid (o_i, o_j) \in L\}$. **Vacuously extending $m^{\Theta_{ij}}$ in \mathcal{L}** yields

$$m^{\Theta_{ij} \uparrow \mathcal{L}}(\mathcal{L}_{ij}) = \alpha_{ij}, \quad m^{\Theta_{ij} \uparrow \mathcal{L}}(\overline{\mathcal{L}_{ij}}) = \beta_{ij}, \quad m^{\Theta_{ij} \uparrow \mathcal{L}}(\mathcal{L}) = 1 - \alpha_{ij} - \beta_{ij}$$

- combining the pairwise mass functions** using Dempster's rule:

$$m^{\mathcal{L}} = \bigoplus_{i < j} m^{\Theta_{ij} \uparrow \mathcal{L}}$$

Plausibility of a linear order

An integer programming problem

- the plausibility of the combination $m^{\mathcal{L}}$ is:

$$pI(L) = \frac{1}{1 - \kappa} \prod_{i < j} (1 - \beta_{ij})^{\ell_{ij}} (1 - \alpha_{ij})^{1 - \ell_{ij}},$$

where $\ell_{ij} = 1$ if $(o_i, o_j) \in L$ and 0 otherwise

- its logarithm $pI(L)$ can be maximized by solving the following **binary integer programming** problem:

$$\max_{\ell_{ij} \in \{0,1\}} \sum_{i < j} \ell_{ij} \ln \left(\frac{1 - \beta_{ij}}{1 - \alpha_{ij}} \right)$$

$$\text{subject to: } \begin{cases} \ell_{ij} + \ell_{jk} - 1 \leq \ell_{ik}, & \forall i < j < k & (1) \\ \ell_{ik} \leq \ell_{ij} + \ell_{jk}, & \forall i < j < k & (2) \end{cases}$$

- constraint (1) ensures that $\ell_{ij} = 1$ and $\ell_{jk} = 1 \Rightarrow \ell_{ik} = 1$, and (2) ensures that $\ell_{ij} = 0$ and $\ell_{jk} = 0 \Rightarrow \ell_{ik} = 0$.

On preference aggregation

A summary

- the framework of belief functions allows us to **model uncertainty in paired comparisons**
- the **most plausible linear order** can be computed efficiently using a binary linear programming approach
- the approach has been applied to **label ranking**, in which the task is to **learn a “ranker”** that maps p -dimensional feature vectors x describing an agent to a linear order over a finite set of alternatives, describing the agent's preferences [Denœux and Masson, 2012]
- the method can easily be extended to the elicitation of preference relations with **indifference** and/or **incomparability** between alternatives [Denœux and Masson. *AOR* 195(1):135-161, 2012]

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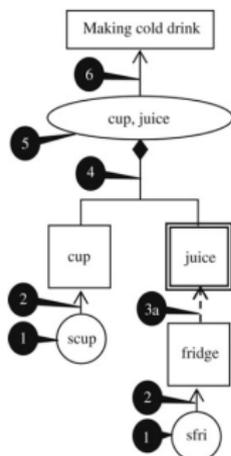
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New applications

of belief functions

- sensor fusion has always been a stronghold of belief calculus
- mainly about merging different sensors using Dempster's rule
- typical applications: tracking and data association, reliability in engineering, image processing, robotics, medical imaging and diagnosis, business and finance (audit)
- a new wave of applications, on:
 - geographical information systems
 - communication networks and security
 - earth sciences
- here we present one (or two!) in more detail:
 - **climate change**
 - **motion capture** in computer vision



(a) Making cold drink.

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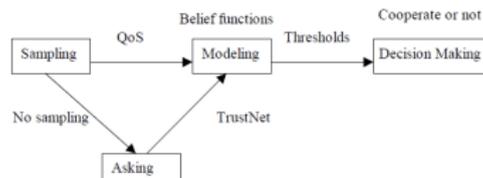
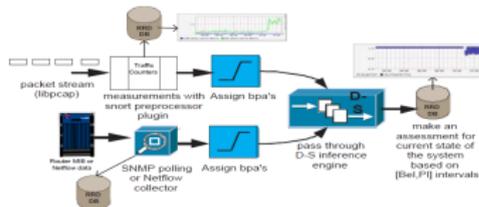
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Most popular applications

of belief functions

- information quality in **financial accounting** [A conceptual framework and belief-function approach to assessing overall information quality (158)]
- auditing** [The Bayesian and belief-function formalisms: A general perspective for auditing (148)]
- reputation and **trust management in telecoms** [An evidential model of distributed reputation management (615)]
- security [An information systems security risk assessment model under the DS theory of belief functions (137)]
- DoS** [Towards multisensor data fusion for DoS detection (137)]



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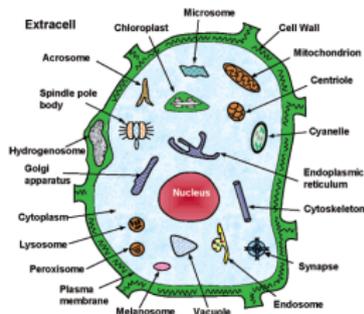


Figure 1. Schematic illustration to show the 22 subcellular locations of eukaryotic proteins: (1) acrosome, (2) cell wall, (3)

- molecular biology** [Ensemble classifier for protein fold pattern recognition (257)], [Predicting eukaryotic protein subcellular location by fusing optimized evidence-theoretic K-nearest neighbor classifiers (205)]

- medical imaging** [Some aspects of Dempster-Shafer evidence theory for classification of multi-modality medical images taking partial volume effect into account (218)]
- earth sciences** and ecology [Integration of geophysical and geological data using evidential belief function (106)], [Decision support system for the sustainable forest management (131)]
- context-aware HCI** and sensing [Sensor fusion using Dempster-Shafer theory (238)], [Evidential fusion of sensor data for activity recognition in smart homes (158)]

Most popular applications

of belief functions

- measurement theory [Measurement uncertainty: An approach via the mathematical theory of evidence (77)]
- **reliability** [Engine fault diagnosis based on multi-sensor information fusion using Dempster-Shafer evidence theory (243)]
- **engineering** [Modelling global risk factors affecting construction cost performance (303)]
- semantic web and **information retrieval** [Dempster-Shafer's theory of evidence applied to structured documents: modelling uncertainty (154)], [EDM: a general framework for data mining based on evidence theory (109)]
- reputation management in e-commerce [Distributed reputation management for electronic commerce (257)]
- **climate change** [Utilizing belief functions for the estimation of future climate change (86)]
- chemistry [Application of belief theory to similarity data fusion for use in analog searching and lead hopping (99)]

Most popular applications

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- **robotics and navigation** [An evidential approach to map-building for autonomous vehicles (229)], [Dempster-Shafer theory for sensor fusion in autonomous mobile robots (192)]
- **tracking and data association** [Shafer-Dempster reasoning with applications to multisensor target identification systems (317)]
- **image processing and computer vision** [Image annotations by combining multiple evidence Wordnet (231)], [Evidence-based recognition of 3-D objects (176)]
- **biometrics** [Image quality assessment for iris biometric (160)]

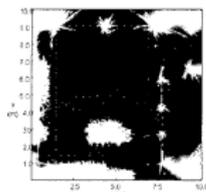
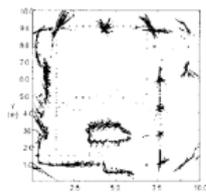


Fig. 10. Occupancy maps generated using the Bayes' fusion rule, values above and below 0.5 separated into full and empty map res. and normalized into the interval $[0, 1]$ for comparison to the D-S method.



108037
TM:people field flowers
cat tiger horses swimmers
TMHD: people field cat
tiger



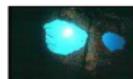
120014
TM:water tree people
buildings light crab
TMHD:people buildings



118033
TM:sky water people sunset
TMHD:sky water people



120070
TM:sky water tree snow cars
TMHD:water cars



12055
TM:sky water tree people
ocean
TMHD:water ocean



121039
TM:water tree people
grass bear stone
TMHD:water tree grass stone

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Climate change

Adaptation of flood defense structures

- climate change is expected to have **enormous economic impact**, including threats to infrastructure assets through
 - damage or destruction from extreme events;
 - coastal flooding and inundation from sea level rise, etc.
- **adaptation of infrastructure** to climate change is a major issue
- engineering design processes and standards are based on **analysis of historical climate data** (using, e.g. Extreme Value Theory), with the assumption of a stable climate
- commonly, flood defenses in coastal areas are designed to withstand at least **100 years return period events**. However, due to climate change, they will be subject during their life time to higher loads than the design estimations
- the main impact is related to the **increase of the mean sea level**, which affects the frequency and intensity of surges
- for adaptation purposes, statistics of extreme sea levels derived from historical data should be combined with projections of the future sea level rise (SLR)

Assumptions and approach

- the **annual maximum sea level** Z at a given location is often assumed to have a Gumbel distribution

$$P(Z \leq z) = \exp \left[- \exp \left(- \frac{z - \mu}{\sigma} \right) \right]$$

with mode μ and scale parameter σ

- current design procedures are based on the **return level** z_T associated with a return period T , defined as the quantile at level $1 - 1/T$: $z_T = \mu - \sigma \log \left[- \log \left(1 - \frac{1}{T} \right) \right]$
- because of climate change, it is assumed that the distribution of annual maximum sea level at the end of the century will be **shifted to the right**, with shift equal to the SLR : $z'_T = z_T + SLR$
- approach**:
 - represent the evidence on z_T by a likelihood-based belief function using past sea level measurements;
 - represent the evidence on SLR by a belief function describing expert opinions;
 - combine these two items of evidence to get a belief function on $z'_T = z_T + SLR$.

Statistical evidence on z_T

- let z_1, \dots, z_n be n i.i.d. observations of Z . The likelihood function is:

$$L(z_T, \mu; z_1, \dots, z_n) = \prod_{i=1}^n f(z_i; z_T, \mu),$$

where the pdf of Z has been reparametrized as a function of z_T and μ

- the corresponding contour function is thus:

$$pl(z_T, \mu; z_1, \dots, z_n) = \frac{L(z_T, \mu; z_1, \dots, z_n)}{\sup_{z_T, \mu} L(z_T, \mu; z_1, \dots, z_n)}$$

and the marginal contour function of z_T is

$$pl(z_T; z_1, \dots, z_n) = \sup_{\mu} pl(z_T, \mu; z_1, \dots, z_n)$$

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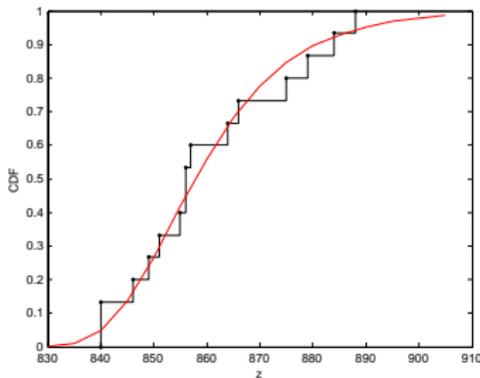
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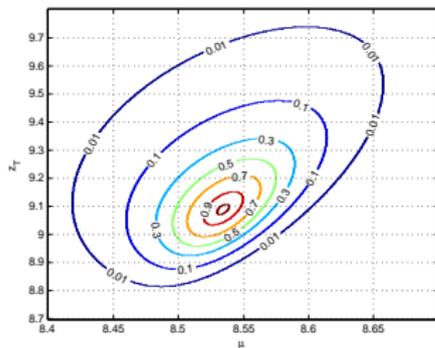
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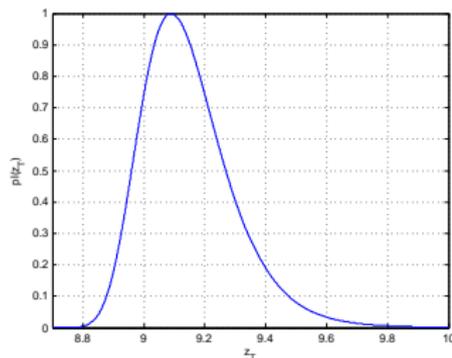
Example of inference



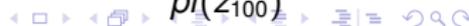
- 15 years of sea level data at Le Havre, France
- corresponding contour function (left) and marginal contour function (right)



$pl(z_{100}, \mu)$



$pl(z_{100})$



Expert evidence on SLR

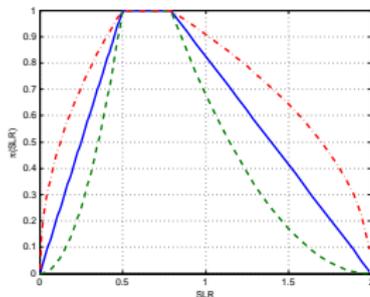
- future SLR projections provided by the IPCC last Assessment Report (2007) give **[0.18 m, 0.79 m]** as a likely range of values for SLR over the 1990-2095 period
- however, it is indicated that **higher values cannot be excluded**
- other recent SLR assessments based on semi-empirical models have been undertaken. For example, based on a simple statistical model, Rahmstorf (2007) suggests **[0.5m, 1.4 m]** as a likely range
- recent studies indicate that **the threshold of 2 m cannot be exceeded** by the end of this century due to physical constraints

Representation

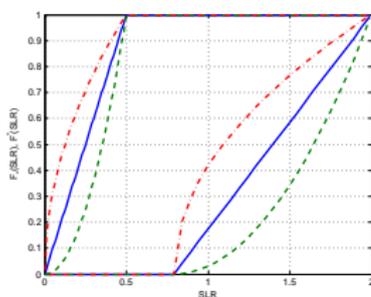
of expert evidence

- the interval $[0.5, 0.79] = [0.18, 0.79] \cap [0.5, 1.4]$ seems to be fully supported, as considered highly plausible by all three sources
- while values outside the interval $[0, 2]$ are considered as impossible
- three representations of expert evidence:
 - consonant random intervals** with core $[0.5, 0.79]$, support $[0, 2]$ and different contour functions π ;
 - p-boxes** with same cumulative *Bel* and *Pl* as above;
 - random sets $[U, V]$ with **independent** U and V and same cumulative belief and plausibility functions as above

Contour functions



Cumulative Bel and Pl



Combination

Principle

- let $[U_{z_T}, V_{z_T}]$ and $[U_{SLR}, V_{SLR}]$ be the **independent random intervals** representing evidence on z_T and SLR , respectively
- the random interval for $z'_T = z_T + SLR$ is

$$[U_{z_T}, V_{z_T}] + [U_{SLR}, V_{SLR}] = [U_{z_T} + U_{SLR}, V_{z_T} + V_{SLR}]$$

- the corresponding belief and plausibility functions are

$$Bel(A) = P([U_{z_T} + U_{SLR}, V_{z_T} + V_{SLR}] \subseteq A)$$

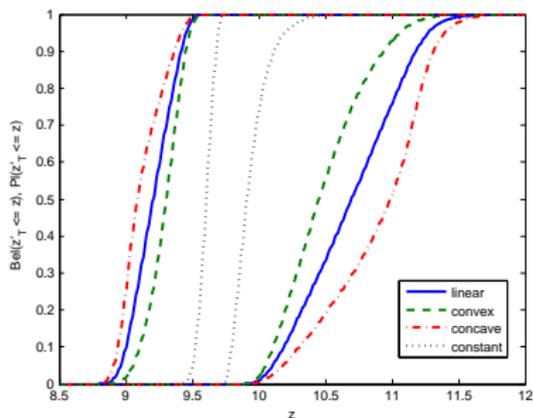
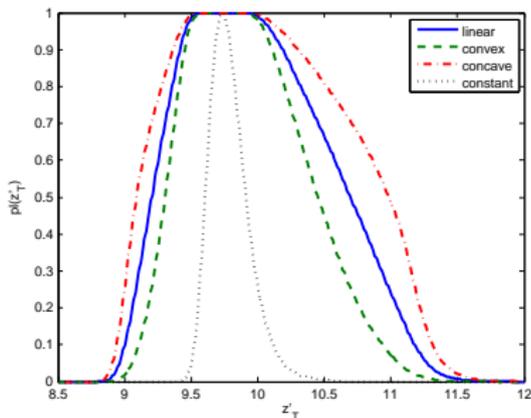
$$Pl(A) = P([U_{z_T} + U_{SLR}, V_{z_T} + V_{SLR}] \cap A \neq \emptyset)$$

for all $A \in \mathcal{B}(\mathbb{R})$

- $Bel(A)$ and $Pl(A)$ can be estimated by **Monte Carlo simulation** (see Computation)

of combining expert and historical belief functions

Results



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Pose estimation

- estimating the position and orientation of an object, along with its internal configuration or pose
- “**model-based**”: explicitly known parametric body model
- “**learning-based**”: exploit the fact that typical (human) motions involve a far smaller set of poses
- directly recover pose estimates **from observable image quantities (features)**

Example-based estimation

- explicitly store a **set of training examples** whose 3D poses are known
- estimate pose by searching for training image(s) similar to the given input image and interpolating from their poses
- **no prior structure** of the pose space is incorporated
- typical architecture:
 - features are extracted from individual images
 - a **map from the features space to the pose space is learned from a training set of examples**
 - the likely pose of the object is then predicted by feeding this feature vector to the learnt map
- approaches: Relevant Vector Machines (RVMs), shape context matching, Local Weighted Regression, BoostMap, Bayesian Mixture of Experts

Scenario

- the available evidence comes in the form of a **training set of images** containing sample poses of an unspecified object
- configuration: a vector $q \in \mathcal{Q} \subset \mathbb{R}^D$
- an **“oracle”** provides for each training image I_k the configuration q_k of the object portrayed in the image
- object location within each training image is known in the form of a **bounding box**
- in **training**, the object explores its range of possible configurations and both samples poses $\tilde{\mathcal{Q}} \doteq \{q_k, k = 1, \dots, T\}$ and N features $\tilde{\mathcal{Y}} \doteq \{y_i(k), k = 1, \dots, T\}, i = 1, \dots, N$ are collected
- source of ground truth: motion capture system
- in **testing**, a supervised localization algorithm is employed to locate the object within the test image
- such features are exploited to produce an estimate of the object's configuration

Pose estimation via belief calculus: why

- let us assume features and poses are described by probability distributions
- as **feature-to-pose maps are typically multi-valued ..**
- ..they induce belief functions on the space of poses
- also, in pose estimation training sets are of limited size
- as in the credal interpretation belief functions amount to a **set of linear constraints on the actual conditional pose distribution (given the features) ..**
- .. they **encode the uncertainty induced by the size of the training set**
- finally, multiple features defined as belief functions on different (feature) domains can be moved to a common refinement (the pose space) and there combined

Training session

- learn from the training data an approximation $\tilde{\rho}$ of the unknown mapping between each feature space \mathcal{Y}_i and the pose space \mathcal{Q}
- we apply **EM clustering** to the N training sequences of feature values $\{y_i(k), k = 1, \dots, T\}, i = 1, \dots, N$..
- ..obtaining a obtain a Mixture of Gaussians (MoG)

$$\left\{ \Gamma_i^j, j = 1, \dots, n_i \right\}, \quad \Gamma_i^j \sim \mathcal{N}(\mu_i^j, \Sigma_i^j)$$

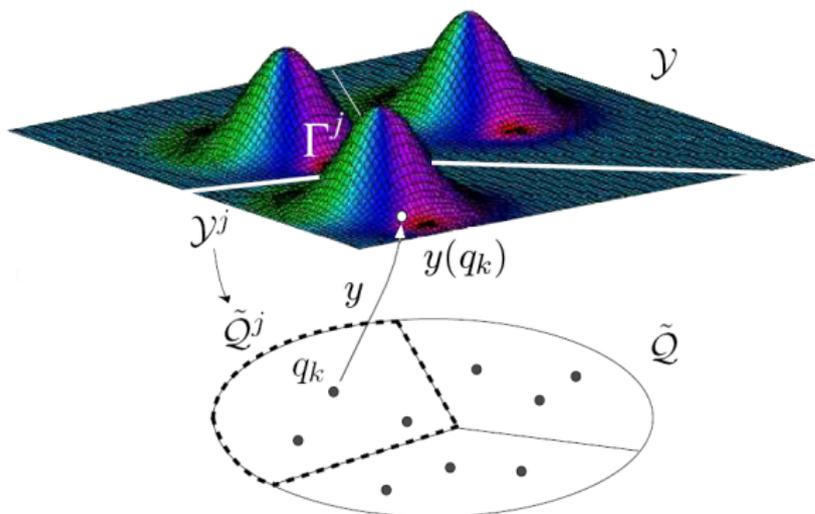
approximation of each feature space \mathcal{Y}_i

- .. and a **discrete approximation of the feature-pose mapping**

$$\rho_i : \mathcal{Y}_i^j \mapsto \tilde{\mathcal{Q}}_i^j \doteq \left\{ q_k \in \tilde{\mathcal{Q}} : y_i(k) \in \mathcal{Y}_i^j \right\}$$

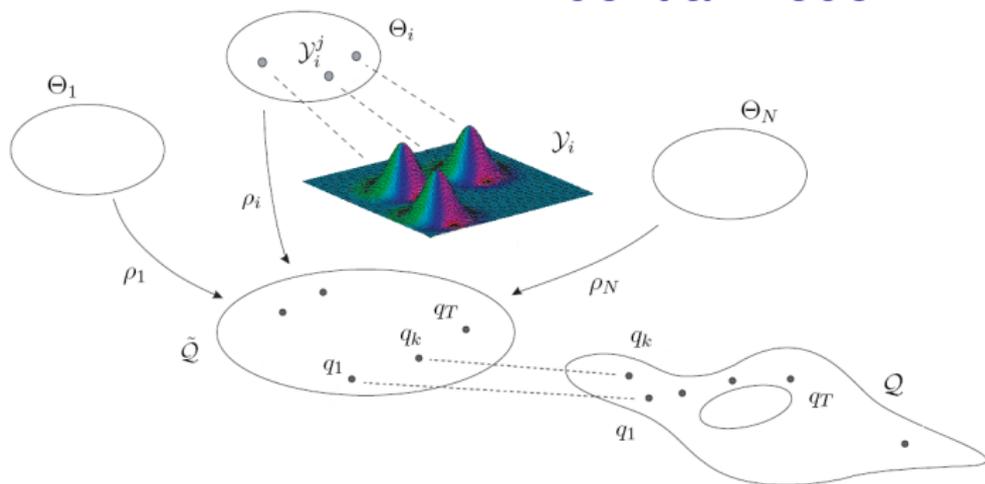
- \mathcal{Y}_i^j is the region of \mathcal{Y}_i in which the j -th Gaussian dominates

Discrete feature-pose maps



- each element \mathcal{Y}_i^j of the approximate feature space is associated with the set of training poses $q_k \in \mathcal{Q}_k$ whose i -th feature value falls in \mathcal{Y}_i^j

Evidential model



- as the applications ρ_i map approximate feature spaces to disjoint partitions of the set of sample poses \tilde{Q} they are refinings
- \tilde{Q} is a common refinement for the approximate feature spaces $\Theta_1, \dots, \Theta_N$
- the collection of FODs $\tilde{Q}, \Theta_1, \dots, \Theta_N$ along with their refinings ρ_1, \dots, ρ_N is characteristic of object to track, features y_i , and training data
- we call it the **evidential model** of the moving object, learned by example

Testing: Dirichlet belief functions

- when one or more new test images are acquired, new visual features y_1, \dots, y_N are extracted
- feature values can be mapped to a collection of belief functions Bel_1, \dots, Bel_N on the set of sample poses \mathcal{Q}
- belief functions also allow to **take into account the scarcity of the training samples..**
- .. by assigning some mass $m(\Theta_i)$ to the whole feature space

$$m_i : 2^{\Theta_i} \rightarrow [0, 1], \quad m_i(\mathcal{Y}_i^j) = \frac{\Gamma_i^j(y_i)}{\sum_k \Gamma_i^k(y_i)} (1 - m_i(\Theta_i))$$

- a reasonable choice is $m_i(\Theta_i) = \frac{1}{n_i}$, as when $n_i \rightarrow \infty$ the discount factor tends to zero and the approximate feature space Θ_i tends to the real feature space \mathcal{Y}_i

Computing pose estimates

Belief estimate

- the belief functions inferred from the test feature values are then mapped to the approximate pose space $\tilde{\mathcal{Q}}$ by **vacuous extension**
- where they are combined by **conjunctive combination**
- this yields the **belief estimate** of the pose, which amounts to a convex set of probabilities on the set of sample poses $\tilde{\mathcal{Q}}$. Then:

- 1 we can compute the **expected pose associated with each vertex** of the credal set:

$$\hat{q} = \sum_{k=1}^T p(q_k) q_k$$

- degree of confidence on the accuracy of the pointwise estimate \rightarrow size of the credal set

- 2 or, we can approximate \hat{b} with a probability \hat{p} on $\tilde{\mathcal{Q}}$ (e.g. the pignistic function)

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Two human pose estimation experiments



- person filmed by two uncalibrated DV cameras
- **arm experiment:** subject moves his arm, while standing in a fixed floor location
- **legs experiment:** person walking normally on the floor, training set collected by sampling a random walk on a section of the floor
- length of the training sequences: 1726 frames for the arm and 1952 frames for the legs
- quite challenging setup: background was highly non-static, with people coming in and out the scene and flickering monitors; self-occlusions

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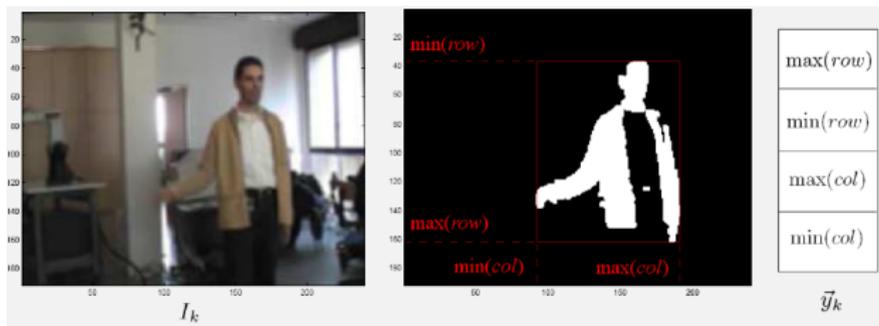
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Annotation and features



- color-based segmentation to get the object of interest (to automatically generate the bounding box annotation required)
- simple feature vector directly from the bounding box: the collection $\max(\text{row})$, $\min(\text{row})$, $\max(\text{col})$, $\min(\text{col})$
- built different evidential models with 2 features from left view, 3 features from right view, or both
- MoG with $n = 5$ components for each feature space

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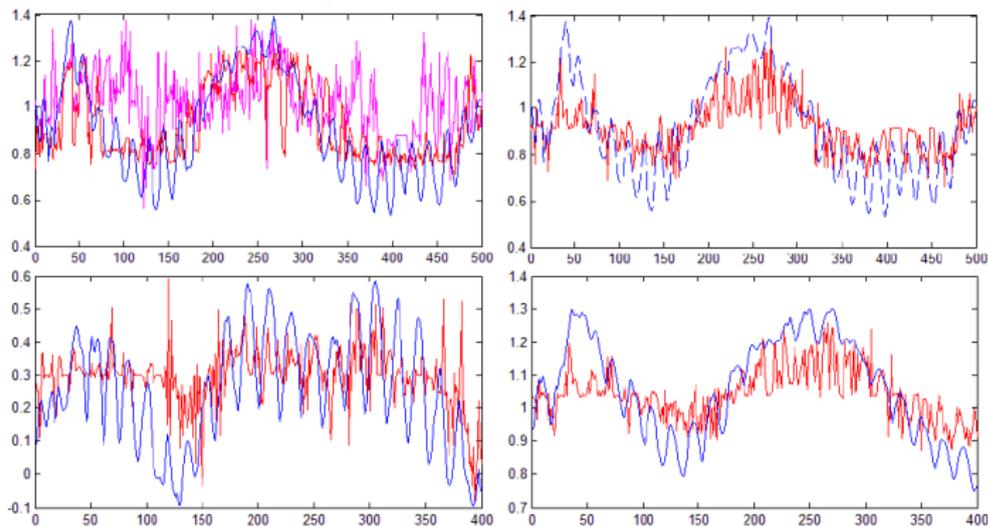
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Empirical results

Comparison with Relevant Vector Machine



- results for component 9 on top, components 1 and 6 at bottom
- blue → ground truth, red → pignistic estimate
- average Euclidean errors: 25.0, 10.6, 18.6, and 7.0 centimeters
- our belief-theoretical approach outperforms the competitor

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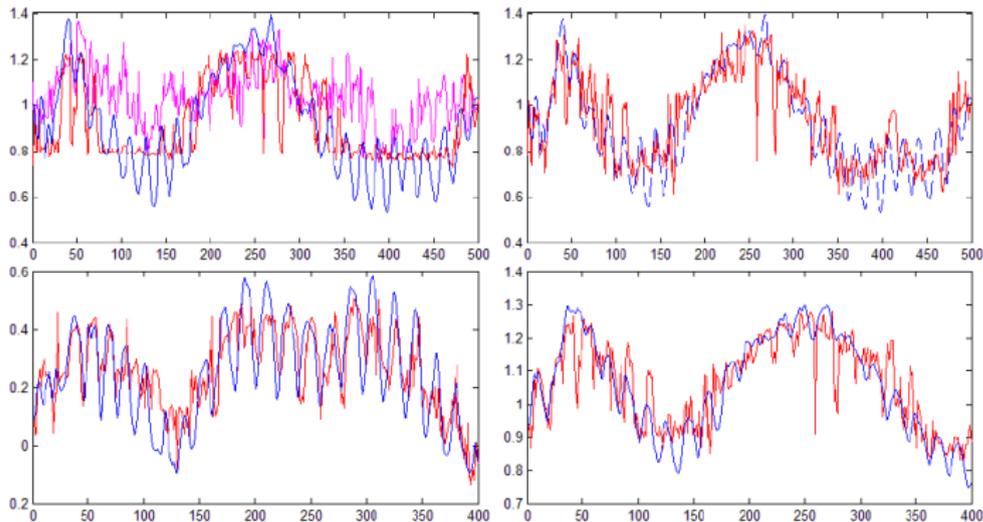
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Empirical results

Comparison with Gaussian Process Regression



- used to build feature-pose maps in, for instance, [14] and [16]
- same components of the pose vector, same test sequence
- our BMR approach clearly outperforms a standard RVM implementation
- average Euclidean errors: 31.2, 13.6, 23.0, and 4.5 centimeters

Conclusions

on the Belief Modeling Regression approach

- presented a novel approach to example-based pose estimation
- framing the problem within belief calculus is natural
- tested in a fairly challenging human pose recovery setup
- exhibits interesting relationships with Gaussian Process approach we did not mention
- future: efficient conflict resolution mechanism
- future: testing of the framework in higher-dimensional pose ranges
- full development of an evidential tracking approach, exploiting temporal information as well via the total belief theorem (see Conditioning)

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A summary

of what we have learned in this tutorial

- the theory of belief functions is a modeling language for **representing elementary items of evidence and combining them**, in order to form a representation of our beliefs about certain aspects of the world
- it is relatively **simple to implement** and has been successfully used in a wide range of applications
- has strong relationships with other theories of uncertainty
- belief functions have interesting mathematical properties in terms of geometry, algebra, combinatorics
- evidential reasoning can be implemented even for **very large spaces and numerous pieces of evidence**, because
 - elementary items of evidence induce **simple belief functions**, which can be combined very efficiently;
 - the **most plausible hypothesis** can be found without computing the whole combined belief function;
 - Monte-Carlo approximations are easily implementable
 - local propagation schemes allow parallelisation

A summary

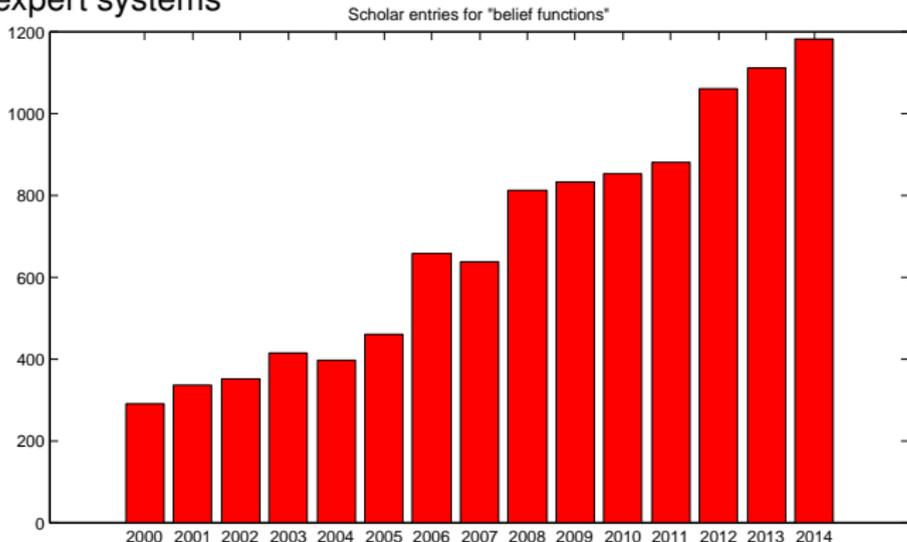
of what we have learned in this tutorial

- statistical evidence may be represented by **likelihood-based belief functions**, generalizing both likelihood-based and Bayesian inference
- inference can also be performed from qualitative data
- decision making strategies based on **intervals of expected utilities** can be formulated that are more cautious than traditional ones
- the **extension to continuous domains** can be tackled via the Borel interval representation, possibly in connection with p-Boxes
- a **toolbox of estimation, classification, regression tools** based on the theory of belief functions is available

Recent trends

in the theory and application of belief functions

- in 2014 alone, almost 1200 papers were published on belief functions and their applications
- new applications are gaining ground, beyond sensor fusion or expert systems





Publications venues

- **conferences** on the theory of uncertainty:

- BFAS's International Conference on Belief Functions (BELIEF)
- Uncertainty in Artificial Intelligence (UAI)
- International Conference on Information Fusion (FUSION)
- International Symposium on Imprecise Probability - Theories and Applications (ISIPTA)
- Symbolic and Quantitative Approaches to Reasoning with Uncertainty (ECSQARU)
- IEEE Systems, Man and Cybernetics (SMC)
- Information Processing and Management under Uncertainty (IPMU)

- **journals** (for theoretical contributions):

- International Journal of Approximate Reasoning (IJAR)
- IEEE Transactions on Fuzzy Systems (I.F. 6.306)
- IEEE Transactions on Cybernetics (I.F. 3.781)
- Artificial Intelligence
- Information Sciences (4.038)
- Fuzzy Sets and Systems

Open issues

and future developments

- what still needs to be resolved:
 - competing epistemic interpretations of belief function theory
 - conditioning issue still open, a variety of approaches proposed depending on semantic adopted and revision principles
 - correct mechanism for evidence combination still debated, depend on meta-information on sources hardly accessible
 - local propagation models (e.g. Shenoy-Shafer) still assume low complexity of local cliques
- what are the next steps?
 - relationships with several fields of mathematics not completely understood
 - generalisation of the total probability theorem in full generality
 - full development of evidential graphical models (e.g. evidential HMMs)
 - tackling current machine learning trends such as transfer learning
 - can it cope with big data paradigm?

For Further Reading

Papers and Matlab software available at:

`https://www.hds.utc.fr/~tdenoeux`

Belief Functions Encyclopedia:

`http://cms.brookes.ac.uk/staff/FabioCuzzolin`

These slides are available online at:

`/FabioCuzzolin/uai-tutorial.pdf`

THANK YOU!

Belief
functions for
the working
scientist

F. Cuzzolin
and
T. Denoeux

Appendix

For Further Reading

For Further Reading I



Visions of a generalized
probability theory

Some original perspectives on the intriguing
mathematics and the practical use of belief
functions



G. Shafer.

A mathematical theory of evidence.

Princeton University Press, 1976.



F. Cuzzolin.

Visions of a generalized probability theory.

Lambert Academic Publishing, 2014.



F. Cuzzolin (Ed.).

Belief functions: theory and applications.

LNCS Volume 8764, Springer, 2014.



For Further Reading I



F. Cuzzolin.

Fifty years of belief functions: a survey. Part I:
Theory

International Journal of Approximate Reasoning (in preparation) 2000.



F. Cuzzolin and C. Sengul.

Fifty years of belief functions: a survey. Part II:
Applications

International Journal of Approximate Reasoning (in preparation) 2000.



F. Cuzzolin.

The geometry of uncertainty.

Springer-Verlag, 2016.